

PETITION ON HORMONAL CONTRACEPTIVES

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Preliminary Statement

Hormonal contraceptives have been on the market for over 50 years and, while their formulations have changed, the basic mechanism of action has remained the same. During this time numerous studies have been performed documenting side effects, some of which appear over time, some within weeks or months, but all can have a serious impact on health. An effort was made to perform a series of comprehensive literature surveys to better understand immediate and long-term side effects of these agents. The results of this literature review have led to several recommendations. These recommendations are listed below with the documentation of the research noted on the following pages.

Action Requested

Drugs which should be removed from the market:

- Depot Medroxyprogesterone Acetate (DMPA)
 - Recommendation to remove from the market the injectable contraceptive Depot Medroxyprogesterone Acetate (DMPA; Depo Provera) based on conclusive evidence that it facilitates the transmission of HIV from men to women. Numerous alternatives are available.

Black box warnings that should be added to prescribing information

- Breast Cancer
 - Combined estrogen-progestogen contraceptives (COCs, including oral, intravaginal and transdermal formulations) are acknowledged by IARC as Group I carcinogens. Substantial data supports an increased risk of breast cancer with the use of COCs. A black box warning should be added to the labeling of all COCs that they have been shown to increase the risk of breast cancer. Patient-related materials should also adequately convey this risk.
 - Progestogen-only contraceptives (POCs) have not been extensively studied, but one large registry study did show a significantly increased risk of breast cancer with use of POCs. Unless there is evidence to the contrary, a similar warning should be added to all POCs. Patient-related materials should also adequately convey this risk.
- Cervical Cancer
 - COCs have been linked to a significantly increased risk of cervical cancer. Similar data have been shown for POCs. A black box warning should be added to the labeling of all COCs and POCs that they have been shown to increase the risk of cervical cancer. Patient-related materials should also adequately convey this risk.
- Inflammatory Bowel Disease
 - Significantly higher risk for the development of inflammatory bowel disease, especially Crohn's disease, but also ulcerative colitis, has been shown for COCs. A black box warning should be added to the labeling of all COCs that their use is linked to a significantly increased risk for the development of inflammatory bowel disease. Patient-related materials should also adequately convey this risk.
- Systemic Lupus Erythematosus (SLE)
 - Significantly higher risk for the development of SLE has been shown for COCs in several studies, especially the best-designed, largest cohort studies. A black box warning should be added to the labeling of all COCs that their use is linked to a significantly increased risk of the development of SLE. Patient-related materials should also adequately convey this risk.
- Depression and Suicide

- Substantive evidence indicates there is a 25% risk of depression for women under 25 years of age especially within 6 months of starting COCs. A black box warning should be added to the labeling of all COCs that their use is linked to a significantly increased risk of the development of depression. Patient-related materials should also adequately convey this risk.
- The relative risk for suicide attempts ranges from 1.91 for COC's, to 2.29 for oral progestins, 2.58 for vaginal ring and 3.28 for patch among adolescents and young women – mean age 21 years – peaking within two months of onset of medication. A black box warning should be added to the labeling of all COCs that their use is linked to a significantly increased risk of suicide. Patient-related materials should also adequately convey this risk. Close monitoring is essential especially in the first year of use.
- Venous Thrombosis and Cardiovascular Events
 - The current black box warning regarding thrombotic events on some formulations, notes “WARNING: CIGARETTE SMOKING AND SERIOUS CARDIOVASCULAR EVENTS.” This is misleading and has shown to be misinterpreted by many women who infer that the increased risk only occurs with cigarette smoking and/or with being over 35 years of age. The warnings should be amended to state, “WARNING: INCREASED RISK OF SERIOUS CARDIOVASCULAR EVENTS INCLUDING BLOOD CLOTS.”
 - This warning should be required for hormonal birth control products including oral, intravaginal and transdermal formulations. The patient-related materials should clearly explain the genetic risk factors, other risk factors, and the signs and symptoms. This warning should be included in ALL direct-to-consumer advertising (television, print, radio, etc.).

Additional safety information which should be added

- Multiple Sclerosis (MS)
 - Significantly higher risk for the development of MS has been shown for COCs in several studies, especially the best-designed, largest case-control studies. A warning should be added to the labeling of all COCs that their use appears to be linked to a significantly increased risk of the development of MS. Patient-related materials should also adequately convey this risk.
- Bone Fractures
 - Use of POCs is clearly associated with a higher risk of bone fractures. A warning should be added to the labeling of all POCs that their use is linked to a significantly increased risk of the development of bone fractures. Patient-related materials should also adequately convey this risk.
 - Protracted use of COCs has been associated with an increased risk of bone fractures. A warning should be added to the labeling of all COCs that their prolonged use may be linked to a significantly increased risk of the development of bone fractures. Patient-related materials should also adequately convey this risk.
- Body Mass Effects
 - For ANY progestin-releasing IUD:
 - Add to professional label in side effects/precautions:
 - Progestin-releasing IUDs (IUCs) have demonstrated in clinical trials to significantly increase % fat body mass with a corresponding decrease in % lean body mass over 1 year of use.
 - Add to patient-related materials:
 - Use of (Brand name) may increase the percent of fat in your body while decreasing the percent of lean body mass; this change in body composition is

known to increase risk of other serious conditions such as diabetes and cardiovascular problems.

- This warning should be included in all direct-to-consumer advertising (television, print, radio, etc.) as it demonstrates use of IUCs may contribute to other serious chronic health conditions.
- Similar labeling should be considered for progestin-only contraceptives. Although the current evidence is less, it tends in the same direction.
- Urogenital Problems
 - Interstitial Cystitis: Significantly higher risk for the development of interstitial cystitis has been shown for COCs in two studies. A warning should be added to the labeling of all COCs that their use appears to be linked to a significantly increased risk of the development of interstitial cystitis. Patient-related materials should also adequately convey this risk.
 - COCs have also been linked to an increased risk of bacteriuria, urinary tract infections, bladder trabeculation, vulvovaginal candidiasis, vaginal dryness, vulvar vestibulitis, and Female Sexual Dysfunction (FSD) caused by OC-induced dyspareunia and reduced sexual desire and libido. These risks should be adequately conveyed in the prescribing information, especially FSD where there is substantial literature evidence.

List of Agents

A list of the agents discussed is shown below. Other than Depot Medroxyprogesterone Acetate (DMPA; Depo Provera) we refer in general to COCs (which refers to all combined estrogen-progestogen contraceptive formulations) and POCs (which refers to all progestin-only contraceptive formulations) regardless of the route of administration (e.g. oral, intravaginal, transdermal, implants, IUS/IUD, etc.).

Combined Estrogen-Progestin (EE-P) Pills

OVCON-35

FEMCON 35

FEMCON FE

BALZIVA 28

BRIELLYN 28

PHILITH

GILDAGIA

VYFEMLA

NEXESTA FE

and generic therapeutic equivalents

BREVICON

MODICON 28

NORMINEST FE

NORTREL 0.5/35-28

WERA

CYCLAFEM

CYONANZ

and generic therapeutic equivalents

GENERESS

KAITLIB FE

and generic therapeutic equivalents

NORINYL 1+35 28-DAY TABLETS
ORTHO-NOVUM 1/35 28 TABLETS
ALYACEN 1/35
ARANELLE
CYCLAFEM 1/35
DASETTA 1/35
NORTREL 1/35-28
NYLIA 1/35
PIRMELLA 1/35

and generic therapeutic equivalents

ORTHO-NOVUM 7/7/7-28
ALYACEN 7/7/7
CYCLAFEM 7/7/7
DASETTA 7/7/7
NORTREL 7/7/7
NYLIA 7/7/7
PIRMELLA 7/7/7

TRI-NORINYL 28-DAY
ARANELLE

NORINYL 1+50 28-DAY

LOESTRIN 21 1/20
LOESTRIN 21 1/20 FE
MINASTRIN 24 FE
TAYTULLA
MIBELAS 24 FE

MICROGESTIN 1/20
MICROGESTIN FE 1/20
JUNEL 1/20
GILDESS 1/20 and GILDESS FE 1/20
LARIN 1/20 and LARIN FE 1/20
BLISOVI 1/20 and BLISOVI FE 1/20
AUROVELA 1/20 and AUROVELA 1/20 FE
HAILEY 1/20 and HAILEY FE 1/20
and generic therapeutic equivalents

LOESTRIN 21 1.5/30
LOESTRIN FE
MICROGESTIN 1.5/30
MICROGESTIN FE
AUROVELA 1.5/30

AUROVELA FE 1.5/30
BLISOVI FE 1.5/30
GILDESS 1.5/30
GILDESS FE 1.5/30
JUNEL 1.5/30
JUNEL FE
LARIN 1.5/30
LARIN FE

ESTROSTEP 21
ESTROSTEP FE
TRI-LEGEST 21
TRI-LEGEST FE
and generic therapeutic equivalents

ZOVIA 1/35E-28
KELNOR
and generic therapeutic equivalents

LOW-OGESTREL-28
CRYSELLE
ELINEST

OGESTREL 0.5/50-28

LoSEASONIQUE
LO SIMPESE
and generic therapeutic equivalents

ALESSE
LEVLITE
LESSINA-28
AVIANE-28
BALCOLTRA
AFIRMELLE
FALMINA
ORSYTHIA
VIENVA
and generic therapeutic equivalents

QUARTETTE—91-DAY
FAYOSIM

SEASONALE
INTROVALE
ALTAVERA
AYUNA

QUASENSE
SETLAKIN
LEVORA 0.15/30-28
KURVELO
PORTIA-28
MARLISSA

SEASONIQUE
ASHLYNA
DAYSEE
JAIMIESS
SIMPESSE
and generic therapeutic equivalents

TRIVORA-28
ENPRESSE-28
LEVONEST
ELIFEMME
MYZILRA
and generic therapeutic equivalents

DESOGEN
EMOQUETTE
ENSKYCE
ISIBLOOM
KALLIGA
and generic therapeutic equivalents

KARIVA
KIMIDESS
VIORELE
PIMTREA
VOLNEA
BEKYEE
and generic therapeutic equivalents

CYCLESSA
VELIVET
and generic therapeutic equivalents

ORTHO-CYCLEN-28
SPRINTEC
PREVIFEM
MONO-LINYAH
ESTARYLLA
MILI
and generic therapeutic equivalents

ORTHO TRICYCLEN 28
TRI-SPRINTEC
TRIPREVIFEM-28
TRI-LINYAH
TRI-ESTARYLLA
TRI-MILI
and generic therapeutic equivalents

ORTHO TRI-CYCLEN LO
TRI-PREVIFEM
TRI LO SPRINTEC
TRI-LO-ESTARYLLA
TRI-LO-MILI
and generic therapeutic equivalents

YAZ
LORYNA
NIKKI
MELAMISA
LO-ZUMANDIMINE
and generic therapeutic equivalents

BEYAZ
and generic therapeutic equivalents

YASMIN 28
SYEDA
YAELA
ZUMANDIMINE
and generic therapeutic equivalents

SAFYRAL

NATAZIA

Combined EE-P Contraceptive Patch
ORTHO EVRA
XULANE

Combined EE-P Contraceptive Ring
NUVARING

Progestin-Only Pills

MICRONOR TABLETS

NOR-QD TABLETS

CAMILA

ERRIN

HEATHER

JENCYCLA

INCASSIA

and generic therapeutic equivalents

Progestin-Only Injectable

DEPO PROVERA

Progestin-Only Implants

JADELLE

NEXPLANON

Progestin-Only IUS/IUD

MIRENA IUS

LILETTA IUD

SKYLA IUD

KYLEENA IUD

Statement of Grounds

Risk of HIV Transmission

One of the most common forms of steroidal contraception is the injectable contraceptive: Depot medroxyprogesterone acetate (DMPA). DMPA is highly effective and requires only quarterly injections, as opposed to daily oral ingestion. As a long-acting type of effective contraceptive, it is not unique, as there are other injectable or implantable contraceptives in wide use, e.g., norethisterone enanthate (NET), as well as other delivery systems such as vaginal rings and patches.

However, evidence began emerging in the 1990s, which has become compelling in recent years, that DMPA is unique among contraceptives in its property of facilitating the transmission of HIV. This dangerous characteristic has been abundantly and unequivocally documented through several lines of evidence which are summarized below:

Epidemiological Evidence

- A. Four meta-analyses (3 reports) were published in 2015. Each used different inclusion criteria and compiled the data on different numbers of studies, yet all 4 came up with essentially the same result of significantly increased risk of male-to-female HIV transmission in women using DMPA (Table 1).

Table 1 – Meta-Analyses Evaluating Risk of HIV Transmission with Depot medroxyprogesterone acetate (DMPA)

| Meta-analysis | # Included studies | Pooled Adj. OR or HR (95% CI) |
|----------------------|---------------------|-------------------------------|
| Ralph et al. 2015 | 10 (longitudinal) | HR 1.40 (1.16–1.69) |
| Morrison et al. 2015 | 18 (longitudinal) | HR 1.50 (1.24–1.83) |
| Brind et al. 2015 | 8 (cross-sectional) | OR 1.41 (1.15–1.73) |
| Brind et al. 2015 | 16 (longitudinal) | HR 1.49 (1.28–1.73) |

- B. Ten primary studies (all longitudinal, published between 2003 and 2014, listed in Table 2 below) were methodologically robust enough to meet the inclusion criteria of all 3 published reviews.

Table 2 – Individual Studies of the Effects of DMPA HIV Transmission

| Study | Yr.(s) of study | Pop. size | Nation and locale | Subject source | Months of follow-up | Follow-up interval (months) | Type of data shown | HR or IRR (95% CI) | Weight (%) |
|-------------------|-----------------|-----------|------------------------------------|--------------------------------------|---------------------|-----------------------------|---------------------|----------------------|------------|
| Crook 2014 | 2005–2009 | 8,663 | S Africa, Uganda, Tanzania, Zambia | Microbicide trial sero-disc. couples | 12 | 1 | Inv. Prob. W'ted HR | 1.45 (1.09–1.93) | 16.39 |
| McCoy 2013 | 2003–2007 | 4,913 | South Africa, Zimbabwe | Diaphragm/gel HIV prev. trial | 24 | 3 | MV HR | 1.22 (0.85–1.76) | 13.20 |
| Morrison 2012 | 2004–2007 | 5,567 | South Africa | General population | 9–24 | 3 | MSM HR | 1.27 (0.93–1.73) | 15.32 |
| Wand 2012 | Not reported | 2,236 | Durban, S. Africa | >90% from microbicide trial | Not reported | 3 | MV HR | 2.02 (1.37–2.99) | 12.22 |
| Heffron 2012 | 2004–2010 | 3,790 | 7 African nations | Sero-discordant couples | 12–24 | 3 | MSM HR | 3.93 (1.38–11.21) | 2.81 |
| Morrison 2007 | 1999–2004 | 6,109 | Uganda, Zimbabwe, Thailand | Family planning clinics | 21.5 | 3 | MSM HR | 1.25 (0.88–1.77) | 13.86 |
| Myer 2007 | 2000–2004 | 4,073 | Cape Town, So. Africa | General population | 24 | 6,6, & 12 | MV IRR | 0.75 (0.33–1.69) | 4.36 |
| Kleinschmidt 2007 | 1999–2001 | 551 | Orange Farm, So. Africa | Family planning clinic | 12 | 3 | MV HR | 0.46 (0.06–3.66) | 0.78 |
| Baeten 2007 | 1993–1997 | 779 | Mombasa, Kenya | CSW | 120 | 1 | MV HR | 1.73 (1.28–2.34) | 15.69 |
| Kiddugavu 2003 | 1994–1999 | 5,117 | Rakai, Uganda | General population | 31 | 10 | IRR, MLR | 0.84 (0.41–1.72) | 5.37 |

Importantly, no consistent association has emerged with regard to oral contraceptives or other injectable or implantable contraceptives and the facilitation of HIV transmission.

Mechanistic Evidence

- A. *In vivo* evidence of increased HIV transmission: Heffron et al. (2012) reported the increased presence of HIV-1 RNA in genital fluids of women using DMPA.
- B. *In vitro* evidence of increased HIV replication at the cellular level: Maritz et al. (2018) reported experimental evidence of increased replication of HIV in human blood monocytes with medroxyprogesterone acetate (MPA).
- C. Experimental evidence of agonistic binding to the glucocorticoid receptor (GR) as the mechanism for DMPA's immunosuppression: over the last 15 years, abundant experimental evidence of cytotoxic and immunosuppressive action of DMPA via its agonistic binding to the GR of human leukocytes has been reported (Schindler 2003; Hapgood and Tomasicchio 2010, Hapgood 2014.) Thus, Huijbregts et al. (2014) reported experimental evidence of immunosuppression of human T cells in vitro by MPA. Tomasicchio et al. (2013) reported experimental evidence of increased human T-cell destruction in vitro via the glucocorticoid receptor (GR) with MPA. Hapgood et al. (2014) reported:
“that MPA, unlike NET and progesterone, represses inflammatory genes in human PBMCs (peripheral blood mononuclear cells) in a dose-dependent manner, via the glucocorticoid receptor (GR), at concentrations within the physiologically relevant range. These and published results collectively suggest that the differential GR activity of MPA versus NET may be a mechanism whereby MPA, unlike NET or progesterone, differentially modulates HIV-1 acquisition and pathogenesis in target cells where the GR is the predominant steroid receptor expressed.”
- D. Evidence of mechanism of MPA action at the gene expression level: experimental evidence of MPA-mediated suppression of inflammatory genes via GR in cultured human cells (Govender 2014) demonstrated the suppression of inflammatory genes in cultured human endocervical cells.

Summary and Conclusions:

DMPA, in contrast to all other steroidal contraceptives, has now conclusively been demonstrated to significantly increase the risk of HIV transmission from infected men to women. The robust epidemiological association has been supported by *in vivo* evidence of increased HIV RNA in the female genital tracts of women using DMPA. Moreover, abundant experimental evidence has shown that MPA, due to its agonistic binding of the GR, specifically represses the innate immune responses of both circulating human leukocytes and endocervical cells and allows for increasing HIV replication. The demonstration in the literature of the chain of causation is therefore compelling.

In the United States, where the availability of a wide range of contraceptive drugs and devices is virtually universal, and where, among these contraceptive choices, one and only one particular method—DMPA—is now known to increase the transmission of an often-fatal viral infection (HIV/AIDS), there can be no justification for such a drug's continued availability in the marketplace. It should be removed from the marketplace by the FDA without further delay.

Risk of HIV Transmission References

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Cancer

Papers were accessed from a PubMed literature review as noted (Williams 2018). Each paper was rated based on the parameters noted in the STROBE statement (von Elm et al. 2007).

Breast Cancer

Breast cancer is the most commonly diagnosed cancer (excluding non-melanoma skin cancers) in women in developed nations, including the U.S., with 1.7 million cases diagnosed worldwide annually. It accounts for 20% of all cancers in women.¹ According to the Surveillance, Epidemiology and End Results (SEER) statistics², it is estimated that there are about 3,418,000 women with invasive breast cancer in the USA as well as over 60,000 cases of in situ cancers. There will be about 266,000 new cases of breast cancer in 2018, accounting for 15.3% of all new cancer cases, with about 41,000 deaths, accounting for 6.7% of all cancer deaths. Nulliparity or late childbearing and high body mass index are risk factors for breast cancer as is exposure to COCs and HRT. Any risk factors that are controllable should be minimized. The data for breast cancer is shown split into cohort studies (Table 3), case control studies (Table 4) and meta-analyses (Table 5).

The carcinogenicity of combined estrogen-progestogen contraceptives was evaluated by IARC working groups initially in 1998 (monograph published in 1999) and again in 2005 (monograph published in 2007). This was most recently updated with studies published through May 2008 (IARC 2012). Since that time, several important studies have been published, most of which are supportive of the IARC classification of COCs as Group I carcinogens and in agreement with the IARC evaluation of specific cancer types. In addition, several important studies have been published evaluating COCs and their cancer risk. In 2002 the National Toxicology Program added steroidal estrogen as a known human carcinogen (Report on Carcinogens, Fourteenth Edition available at <https://ntp.niehs.nih.gov/pubhealth/roc/index-1.html>).

In agreement with IARC the recent data confirms an increased risk of breast cancer with use of COCs (Table 2). After 2005, there continue to be studies demonstrating the significant risk of breast cancer with hormonal contraception. In January 2006, the *New England Journal of Medicine* published a review article which found estrogen-progestin drugs increased breast cancer risk (Yager 2006). In October 2006 the *Mayo Clinic Proceedings* published a meta-analysis confirming estrogen-progestin drugs increase premenopausal breast cancer (Kahlenborn 2006).

The studies that looked at recent use (within 1–5 years) or current use of COCs in premenopausal women showed the most dramatic increased risk for breast cancer. In a case control study, women ages 20-49 years with use of COCs within a year had an increased risk of breast cancer (OR, 1.5; 95% CI, 1.3–1.9) (Beaber 2014). The same study showed an increase in risk depending on the formulation with triphasic COCs carrying a markedly increased risk (OR, 3.1; 95% CI, 1.4–4.7). In another large case control study of women ages 20-45 years, use of COCs for a year or more resulted in a 2.5-fold increased risk of triple-negative breast cancer (95% CI 1.4–4.3) but not for the receptor-positive breast cancers. In the same study, women 40 years or younger with a year or more use of COCs had a higher relative risk of triple-negative breast cancer (RR, 4.2; 95% CI, 1.9–9.3) (Dolle 2009). A cohort study of over 35,000 postmenopausal women found a significantly increased risk of breast cancer in women on hormone replacement therapy (HRT) if they had used COCs in the past (RR, 2.45; 95% CI, 1.92–3.12) as compared with never users (RR, 1.67; 95% CI, 1.32–2.12) (Lund 2007). There also appears to be an increased risk for African American women on COCs within the past five years for ER+ cancers (OR, 1.46, 95% CI, 1.18–1.81), for ER- cancers (OR, 1.57; 95% CI, 1.22–1.43) and for triple-negative

¹ <https://www.wcrf.org/int/cancer-facts-figures/data...cancers/breast-cancer-statistics>.

² <https://seer.cancer.gov/statfacts/html/breast.html>.

cancers (OR, 1.78; 95% CI, 1.25–2.53) with the risk of ER+ cancers continuing for 15–19 years after stopping the COCs (Bethea 2015).

In a French study (DeLort 2007) of 934 women who developed breast cancer, the use of COCs increased the risk of early development of breast cancer (OR, 1.84; 95% CI, 1.38–2.44). However, initiating COCs after age 23 reduced the risk (OR, 0.52; 95% CI, 0.34–0.79). Use of the levonorgestrel-releasing IUD, commonly used to treat abnormal bleeding in the perimenopause, increased the risk of developing breast cancer in postmenopausal women (OR, 1.48, 95% CI, 1.10–1.99) (Heikkinen 2016). The risk varies with the formulation as current use of a triphasic pill containing levonorgestrel carries an excess risk of causing breast cancer (RR, 3.05; 95% CI, 2.00–4.66) (Hunter 2010). In a large prospective cohort study of 1.8 million Danish women ages 15 to 49, enrolled and followed from 1995 to 2012 through various national registries, the risk of breast cancer among current or recent users increased depending on length of use from RR, 1.09 with less than one year of use (95% CI, 0.96–1.23) to an RR, 1.38 (95% CI, 1.26–1.51) for more than 10 years of use (Mørch 2017). They found the increased risk persisted after discontinuing use if COCs were used for 5 years or more. These investigators also found an increased risk in current or recent use of the progestogen-only intrauterine device (RR, 1.21; 95% CI, 1.11–1.33).

In most Western countries, 5% to 10% of all breast cancer cases are due to a main genetic cause: mutations of the BRCA1 and BRCA2 genes constitute 90% of hereditary breast cancer cases (Mehrgou 2016). These women are often begun on COCs at an early age to reduce their risk of ovarian cancer. However, in a case control study of 2,492 matched pairs of women with the *BRCA1* gene, COC use was associated with an increased risk of early onset breast cancer if begun under the age of 20 (OR, 1.45; 95% CI, 1.20–1.75) (Kotsopoulos 2014) and the risk increased by 11% for each additional year of use.

More recent publications include data from some very recent, large cohort studies (Mørch 2017, Heikkinen 2016, Poosari 2014) with RRs ranging from 1.2 to 1.37. Since breast cancer is by far the most common cancer in women, affecting 1 in 8 women at some time during their lives, this translates into a substantial number of additional cancer cases. In addition, a large registry study of POCs (Soini 2014) also showed an increased RR for breast cancer of 1.19. Increased duration of use also increases the risk of breast cancer for COCs as does use early in life (Mørch 2017).

Table 3 – Breast Cancer (Cohort Studies)

| Study | Study Design | OR ¹ Ever Use | RR ² Ever Use | OR Current Use | RR Current Use | OR Past Use | RR Past Use | Cases | Controls | Quality Score |
|---------------------------------|----------------|-----------------------------|-----------------------------------|-------------------|---------------------|----------------|----------------|-----------|----------------|---------------|
| Mørch et al. 2017 | Cohort | | 1.2 ³ (1.14–1.26) | | | | | 1,797,932 | * ⁴ | 100% |
| Heikkinen et al. 2016 | Cohort | | 1.37 (1.12–1.68) | | | | | 7,000 | 20,000 | 100% |
| Lund et al. 2007 | Cohort | | 1.33 (1.11–1.59) | | | | | 11,777 | 23,676 | 96% |
| Poosari et al. 2014 | Cohort | | 1.31 (0.65–2.65) | | | | | 70 | 11,344 | 92% |
| Phipps et al. 2011 | * ⁵ | | 0.80 ⁶ (0.68–0.94) | | | | | 5,194 | | 92% |
| Brohet et al. 2007 ⁷ | Cohort | | 1.47 (1.16–1.87) | | | | | 846 | 747 | 88% |
| Thorbjarnardottir et al. 2014 | Cohort | | 1.32 (1.02–1.70) | | | | | 654 | 16,928 | 84% |
| Samson et al. 2017 | Cohort | | 1.80 ⁸ (1.29–2.55) | | | | | 4816 | | 83% |
| Rosenberg et al. 2010 | Cohort | | 1.65 (1.19–2.30) | | | | | 789 | 53,848 | 83% |
| Silvera et al. 2005 | Cohort | | 0.88 ⁹ (0.73–1.07) | | | | | 1,707 | 25,611 | 78% |
| Hunter et al. 2010 | Cohort | | 1.12 (0.95–1.33) | | 1.33 (1.03–1.73) | | | 1,344 | 115,264 | 73% |
| | | | 1.42 ¹⁰ (1.05–1.94) | | | | | | | |

¹ OR = odds ratio (95 % confidence interval).

² RR = relative risk (95 % confidence interval).

³ Initiation before age 20, greater than 10 years of use and evaluation within 5 yrs. of stopping further increased the risk.

⁴ Entire population of Denmark was the cohort.

⁵ Concurrent randomized clinical trials and an observational study.

⁶ Hazard ratio shown. Note that women started COCs after age 25, had been off COCs for many years.

⁷ Evaluation in patients carrying BRCA mutations. Hazard ratios shown.

⁸ Hazard ratio shown.

⁹ Hazard ratio shown.

¹⁰ Eight or more years of use.

| Study | Study Design | OR ¹ Ever Use | RR ² Ever Use | OR Current Use | RR Current Use | OR Past Use | RR Past Use | Cases | Controls | Quality Score |
|-----------------------------------|--------------|-----------------------------|-----------------------------------|---------------------|-------------------|----------------|----------------|-------------------|---------------------|---------------|
| | | | 3.05 ¹¹ (2.00–4.66) | | | | | | | |
| Trivers et al. 2007 ¹² | Cohort | | | 1.57 (0.95–2.61) | | | | 292 ¹³ | 1,264 ¹⁴ | 67% |

¹¹ Levonorgestrel containing combined oral contraceptives.

¹² Looked at mortality in patients with breast cancer over 8-10 years depending on whether they were on COCs at the time of diagnosis or within one year.

¹³ Deaths.

¹⁴ Total cohort.

Table 4 – Breast Cancer (Case Control Studies)

| Study | Study Design | OR ¹⁵ Ever Use | RR ¹⁶ Ever Use | OR Current Use | RR Current Use | OR Past Use | RR Past Use | Cases | Controls | Quality Score |
|---------------------------------------|-------------------------|-----------------------------------|------------------------------|--------------------------------|-------------------|---------------------|----------------|-------|----------|---------------|
| Dolle et al. 2009 | Case control | 2.5 (0.9-5.24) | | 4.2 (1.9-9.3) | | | | 898 | 961 | 100% |
| Lee et al. 2008 | Case Case ¹⁷ | 0.68 (0.33-1.38) | | | | | | 94 | 444 | 100% |
| Sweeney et al. 2007 | Case control | 1.27 (0.99-1.63) | | | | | | 2,318 | 2,515 | 100% |
| Beaber et al. 2014b | Case control | 1.5 (1.1-2.2) | | | | | | 985 | 882 | 100% |
| Li et al. 2012 ¹⁸ | Case control | 2.2 (1.2-4.2) | | | | | | 1,028 | 919 | 96% |
| Beaber et al. 2014a | Case control | | | 1.5 ¹⁹ (1.3-1.9) | | | | 1,102 | 21,952 | 96% |
| Ichida et al. 2015 | Case control | | | 0.45 (0.22-0.90) | | | | 155 | 12,333 | 96% |
| Ma et al. 2010 | Case control | 2.87 ²⁰ (1.44-5.74) | | | | | | 1,197 | 2,015 | 95% |
| Folger et al. 2007 | Case control | 1.0 ²¹ (0.8-1.1) | | | | | | 4575 | 4682 | 92% |
| Jernstrom et al. 2005 | Case control | | | | | 2.10 (1.32-3.33) | | 245 | 745 | 92% |
| Kotsopoulos et al. 2014 ²² | Case control | 1.45 ²³ (1.20-1.75) | | | | | | 2,492 | 2,492 | 88% |
| | | 1.19 ²⁴ (0.99-1.42) | | | | | | | | |
| Figueiredo et al. 2010 ²⁵ | Case control | | | | | 2.38 (0.72-7.83) | | 705 | 1,398 | 86% |

¹⁵ OR = odds ratio (95 % confidence interval).

¹⁶ RR = relative risk (95 % confidence interval).

¹⁷ BRCA1 and BRCA2 carriers with breast cancer.

¹⁸ Population-based case-control of women 20-44 yo with recent DMPA use for at least 12 months.

¹⁹ Use within the past year of COCs increases risk of breast cancer.

²⁰ Triple negative breast cancer if less than 18 yo on COCs.

²¹ Evaluated short-term use only.

²² Study of BRCA+ patients.

²³ <20 years old.

²⁴ 20-25 years old.

²⁵ Evaluation of BRCA1 and BRCA2 carriers; controls with unilateral breast cancer compared with contralateral cases.

| Study | Study Design | OR ¹⁵ Ever Use | RR ¹⁶ Ever Use | OR Current Use | RR Current Use | OR Past Use | RR Past Use | Cases | Controls | Quality Score |
|----------------------|-------------------------|------------------------------|------------------------------|-------------------|-------------------|----------------|----------------|-------|----------|---------------|
| Veneroso et al. 2008 | Case Case ²⁶ | 1.12 (1.03-1.23) | | | | | | 116 | 99 | 86% |

| Study | Study Design | OR ²⁷ Ever Use | RR ²⁸ Ever Use | OR Current Use | RR Current Use | OR Past Use | RR Past Use | Cases | Controls | Quality Score |
|----------------------------|--------------------------------|-----------------------------------|------------------------------|-------------------|-------------------|---------------------|----------------|-------|----------|---------------|
| Ma et al. 2006 | Case control | 1.27 ²⁹ (0.75-2.14) | | | | 0.76 (0.49-1.18) | | 1,366 | 440 | 84% |
| | | 0.76 ³⁰ (0.49-1.18) | | | | | | | | |
| Rosenberg et al. 2008 | Case control | 1.5 ³¹ (1.2-1.8) | | | | | | 907 | 1,711 | 83% |
| Haile et al. 2006 | Case control | 0.77 ³² (0.53-1.12) | | | | | | 195 | 497 | 83% |
| | | 1.62 ³³ (0.90-2.92) | | | | | | 128 | 307 | |
| Milne et al. 2005 | Case control | 1.52 (1.22-1.91) | | | | | | 1156 | 815 | 83% |
| Amadou et al. 2013 | Case control | 1.68 (0.67-4.21) | | | | | | 1,000 | 1,074 | 75% |
| Ozmen et al. 2009 | Case control | 0.60 (0.48-0.74) | | | | | | 1,492 | 2,167 | 74% |
| Delort et al. 2007 | Population based ³⁴ | 1.84 ³⁵ (1.38-2.44) | | | | | | 934 | | 71% |
| Beji et al. 2006 | Case control | 1.98 (1.38-2.85) | | | | | | 405 | 1,050 | 63% |
| Veisy et al. 2015 | Case control | 2.11 (1.44-3.08) | | | | | | 235 | 235 | 63% |
| Adams-Campbell et al. 2010 | Case control | 2.83 (1.87-4.24) | | | | | | 321 | 321 | 58% |

²⁶ Comparison of more aggressive with less aggressive cases.

²⁷ OR = odds ratio (95 % confidence interval).

²⁸ RR = relative risk (95 % confidence interval).

²⁹ ER-/PR-

³⁰ ER+/PR+

³¹ OR for 5+ years of use.

³² BRCA1+ patients.

³³ BRCA2+ patients.

³⁴ Population-based study of early onset breast cancer.

³⁵ OR for developing breast cancer 2 years earlier than non-users.

| Study | Study Design | OR ²⁷ Ever Use | RR ²⁸ Ever Use | OR Current Use | RR Current Use | OR Past Use | RR Past Use | Cases | Controls | Quality Score |
|---------------------|----------------------|------------------------------|------------------------------|-------------------|-------------------|----------------|----------------|-------|----------|---------------|
| Lumachi et al. 2010 | Retrospective Review | 2.06 (1.14-3.70) | | | | | | 404 | 408 | 33% |

Table 5 – Breast Cancer (Meta-Analyses)

| Study | Study Design | OR ³⁶ Ever Use | RR ³⁷ Ever Use | OR Current Use | RR Current Use | OR Past Use | RR Past Use | Cases | Controls | Quality Score |
|--------------------------------------|---------------|-----------------------------------|------------------------------|-------------------|-------------------|----------------|----------------|--------|----------|---------------|
| Kahlenborn et al. 2006 ³⁸ | Meta-analysis | 1.19 (1.09-1.29) | | | | | | 18,406 | 27,677 | 91% |
| | | 1.29 ³⁹ (1.20-1.40) | | | | | | | | |
| | | 1.24 ⁴⁰ (0.92-1.67) | | | | | | | | |
| | | 1.44 ⁴¹ (1.28-1.62) | | | | | | | | |
| Bethea et al. 2015 | Meta-analysis | 1.46 ⁴² (1.18-1.81) | | | | | | 1,848 | 10,044 | 85% |
| | | 1.57 ⁴³ (1.22-1.43) | | | | | | 1,043 | 10,044 | |
| | | 1.78 ⁴⁴ (1.25-2.53) | | | | | | 494 | 10,044 | |
| Zhu et al. 2012 | Meta-analysis | 1.08 ⁴⁵ (0.99-1.17) | | | | | | | | 54% |
| Friebel et al. 2014 ⁴⁶ | Meta-analysis | 1.36 ⁴⁷ (0.99-1.88) | | | | | | | | 27% |
| | | 1.51 ⁴⁸ (1.10-2.08) | | | | | | | | |
| Moorman et al. 2013 | Meta-analysis | 1.21 ⁴⁹ (0.93-1.58) | | | | | | | | |

³⁶ OR = odds ratio (95 % confidence interval).

³⁷ RR = relative risk (95 % confidence interval).

³⁸ Limited to case-control studies from 1980-2004.

³⁹ Parous women.

⁴⁰ Nulliparous women.

⁴¹ Use before first full term pregnancy among parous women.

⁴² ER+

⁴³ ER-

⁴⁴ Triple negative.

⁴⁵ For each 5 years on COCs the risk increased by 7%, but statistical significance not achieved.

⁴⁶ Study limited to BRCA1 and BRCA2 mutation carriers.

⁴⁷ 1-3 years of use.

⁴⁸ >3 years of use.

⁴⁹ 8 studies on BRCA1+ or BRCA2+ patients and breast cancer risk with CSC use.

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Cervical Cancer

According to the SEER statistics¹, it is estimated that there are 257,524 women in the US with cervical cancer. There will be about 13,000 new cases of cervical cancer in 2018, with about 4,000 deaths. The five-year survival for cervical cancer is 66%. The IARC evaluation of an increased risk of cervical cancer with COCs is also supported especially by a large, high-quality cohort study (Roura 2016, Table 6). The data for cervical cancer presented in Table 4 shows in particular a higher risk for invasive cervical cancer, and a higher risk with current use. All studies appear to agree that there is an increased risk of cervical cancer in users of COCs (OR apparently about 1.05 per year of use), and this risk increases with duration of use. Current use appears to confer a higher risk than past use, and the risk for invasive cancer shows the highest increase in risk (Roura 2016). A meta-analysis of case-control studies that focused on patients positive for human papilloma virus DNA (Moreno 2002) also showed an increased risk, especially with protracted (5+ years) of use of COCs. One case-control study (McFarlane-Anderson 2008) and one meta-analysis (International Collaboration 2007) also showed an increased risk with progestogen-only contraceptives. Thus, there does appear to be an increased risk of cervical cancer in users of COCs or POCs, and the risk appears to increase with duration of use.

¹ <https://seer.cancer.gov/statfacts/html/corp.html>

Table 6 – Cervical Cancer

| Study | Study Design | OR Ever Use | RR Ever Use | OR Current Use | RR Current Use | OR Past Use | RR Past Use | Cases | Controls | Quality Score |
|---|--------------------------------|----------------------------------|-------------------------------|----------------|--------------------------------|-------------|-------------------------------|--------|----------|---------------|
| Roura et al. 2016 | Cohort Study | | 1.1 ¹ (0.9–1.3) | | 1.8 ¹⁰ (1.4–2.4) | | 1 ¹⁰ (0.9–1.3) | 1,065 | 306,971 | 94% |
| | | | 1.6 ² (1.1–2.3) | | 2.2 ⁸ (1.3–4.0) | | 1.6 ⁸ (1.1–2.2) | 261 | 306,971 | |
| Leslie et al. 2014 | Case Control Study | 1.35 ³ (0.99-1.85) | | | | | | 219 | 2,300 | 87% |
| McFarlane-Anderson et al. 2008 | Case Control Study | 1.59 ⁴ (0.87-2.82) | | | | | | 240 | 102 | 83% |
| | | 2.48 ⁵ (1.30-4.74) | | | | | | | | |
| Vanakankovit et al. 2008 | Case Control Study | 1.49 (0.79-2.64) | | | | | | 60 | 180 | 76% |
| Wilson et al. 2013 | Case Control Study | 1.22 (0.96–1.56) | | | | | | 724 | 3,479 | 76% |
| Matos et al. 2005 | Case Control Study | 1.3 (0.8–3.1) | | | | | | 140 | 157 | 47% |
| International Collaboration 2007 ⁶ | Meta-analysis | 1.057 (1.04–1.07) | | | | | | 16,573 | 35,509 | 97% |
| | <5 years of use | 0.96 (0.04)8 | | | | | | | | |
| | 5-9 years of use | 1.2 (0.05)5 | | | | | | | | |
| | 10+ years of use | 1.56 (0.08)5 | | | | | | | | |
| | <5 years of use | 1.07 (0.08)9 | | | | | | 7,227 | 19,335 | |
| | 5+ years of use | 1.22 (0.11)6 | | | | | | | | |
| Moreno 2002 ¹⁰ | Meta-analysis | | | | | | | 1676 | 255 | 95% |
| | Invasive cervical cancer (ICC) | 1.29 (0.88-1.91) | | | | | | | | |
| | ICC 5+ years of use | 4.01 (2.01-8.02) | | | | | | | | |

¹ Includes Cervical Intraepithelial Neoplasia Grade 3, carcinoma in situ and invasive cervical cancer.

² Analysis limited to invasive cervical cancer.

³ Study limited to HIV+ women.

⁴ Combined hormonal contraceptives.

⁵ Progesterone only contraceptives.

⁶ Meta-analysis of 24 studies (15 cohort and 9 case-control studies).

⁷ Relative risk per year of use for current users of combined hormonal contraceptives.

⁸ Floating standard error shown for users of combined hormonal contraceptives.

⁹ Progestin only contraceptives. Floating standard error shown. The 95% CI for 5+ years of use is 1.01-1.46.

¹⁰ Pooled data from 8 case-control studies of invasive cervical cancer and 2 of carcinoma in situ, analyzing only the subset positive for Human Papilloma Virus DNA in cervical cells.

| Study | Study Design | OR Ever Use | RR Ever Use | OR Current Use | RR Current Use | OR Past Use | RR Past Use | Cases | Controls | Quality Score |
|-------|-------------------------|---------------------|----------------|-------------------|-------------------|----------------|----------------|-------|----------|------------------|
| | In situ carcinoma (ISC) | 1.42 (0.99-2.04) | | | | | | | | |
| | ISC 5+ years of use | 3.42 (2.13-5.48) | | | | | | | | |

Cervical Cancer References

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Crohn's Disease

Papers were accessed from a PubMed literature review as noted (Williams 2017). Each paper was rated based on the parameters noted in the STROBE statement (von Elm 2007).

Overall, 17 primary studies and two meta-analyses were identified which evaluated the effect of COCs on the later development of Crohn's disease (Table 7). Of the 17 primary studies, 4 showed a significantly increased risk for either ever use (Ng 2012, Sicilia 2001, Katschinski 1993) or current use (Katschinski 1993, Khalili 2013) or past use (Khalili 2013). None of the primary studies showed a significantly decreased risk. One meta-analysis (Godet 1995) gave a significantly increased RR of 1.44 (95% CI 1.12–1.86) for ever use of COCs. A meta-analysis published in 2008 showed a significantly increased risk for current use (RR of 1.46 [1.26–1.70]) compared with 1.04 (0.816–1.340) for past use. Recent studies have produced similar findings as older studies, with the highest OR published in 2012 (9.04 [1.11–73.6]). Overall these studies indicate that use of COCs conveys an increased risk of Crohn's disease, especially current use.

Table 7 – Individual Studies of the Effects of COCs on the Development of Crohn’s Disease

| Study | Study Design | OR Ever Use | RR Ever Use | OR Current Use | RR Current Use | OR Past Use | RR Past Use | Cases | Controls | Quality Score |
|---|--------------|------------------|---------------------|---------------------|---------------------|--------------------|---------------------|--------|----------|---------------|
| Khalili et al. 2013 ¹ | Cohort | | 1.43 | | 2.82 (1.65–4.82) | | 1.39 (1.05–1.85) | 315 | 117,060 | 93% |
| García Rodríguez et al. 2005 ² | Cohort | | | | 1.94 (0.85–4.45) | | 1.04 (0.50–2.17) | 171 | 10,000 | 88% |
| Logan and Kay 1989 | Cohort | | 1.7 (0.88–3.2) | | | | | 42 | 45,958 | 54% |
| Vessey et al. 1986 ³ | Cohort | | | | 1.33 | | | 18 | 17,014 | 46% |
| Boyko et al. 1994 | Case-control | | 2 (1.0–3.7) | | | | | 91 | 169 | 94% |
| Katschinski 1993 ⁴ | Case-control | | | | 2.5 (0.75–4.6) | | | | | 93% |
| Katschinski 1993 ⁵ | Case-control | | | | 3.1 (1.1–6.7) | | | | | 93% |
| Lashner et al. 1989 | Case-control | 1 (0.46–2.16) | | 0.73 (0.34–1.59) | | 1.8 (0.61–5.29) | | 51 | 51 | 88% |
| Lesko et al. 1985 ⁶ | Case-control | | 1.7 (1.0–3.2) | | | | | 57 | 2189 | 83% |
| Sandler et al. 1992 | Case-control | | 1.49 (0.99–2.26) | | | | | 184 | 217 | 81% |
| Persson et al. 1993 | Case-control | | 1.7 (0.9–3.2) | | | | | 152 | 305 | 81% |
| Halfvarson et al. 2006 ⁷ | Case-control | | | | 1.5 (0.4–5.3) | | | 102 | 102 | 75% |
| Lowe et al. 2009 ⁸ | Case-control | | 1.05 | | | | | 21,172 | 754,6131 | 74% |
| Ng et al. 2012 ⁹ | Case-control | 4 (1.1–14.2) | | | | | | 125 | 125 | 74% |

¹ Hazard ratios (RR adjusted for time).

² OR increased with duration of use.

³ Authors’ calculation adjusted for smoking.

⁴ Adjusted RR for 1-3 years prior to disease onset.

⁵ Adjusted RR for >3 years prior to disease onset.

⁶ RR is from multiple logistic regression analysis.

⁷ Monozygotic and dizygotic twins.

⁸ Adjusted incidence rate ratio.

⁹ Twins study.

| Study | Study Design | OR Ever Use | RR Ever Use | OR Current Use | RR Current Use | OR Past Use | RR Past Use | Cases | Controls | Quality Score |
|-----------------------------------|--------------------------|---------------------|---------------------|-------------------|---------------------|------------------|----------------------|-------|----------|------------------|
| Ng et al. 2012 ¹⁰ | Case-control | 9.04 (1.11–73.6) | | | | | | | | 74% |
| Sicilia et al. 2001 | Case-control | 2.8 (1.01–7.77) | | | | | | 103 | 103 | 71% |
| Corrao et al. 1998 | Case-control ever use | | | 3.4 (1.0–11.9) | | 1.8 (0.4–7.3) | | 225 | 225 | 67% |
| Katschinski 1993 ¹¹ | Case-control | | 4.3 (1.3–14.4) | | | | | 83 | 83 | 57% |
| Han et al. 2010 | Case-control | | 0.66 (0.38–1.15) | | | | | 315 | 536 | 52% |
| Calkins et al. 1986 ¹² | Case-control | 1.14 (0.44–2.96) | | | | | | 66 | 67 | 42% |
| Calkins et al. 1986 ¹³ | Case-control | 1.6 (0.59–4.37) | | | | | | 66 | 71 | 42% |
| Vcev et al. 2015 | Case-control | 0.28 (0.03–2.46) | | | | | | 11 | 42 | 31% |
| Cornish et al. 2008 | Meta-analysis | | | | 1.46 (1.26–1.70) | | 1.04 (0.816–.340) | 1251 | 74,564 | 91% |
| Cornish et al. 2008 ¹⁴ | Meta-analysis | | | | 1.58 (1.07–2.40) | | | | | 91% |
| Godet et al. 1995 ¹⁵ | Meta-analysis | | 1.44 (1.12–1.86) | | | | | 531 | 49,156 | 82% |

¹⁰ Multivariate analysis.

¹¹ RR for use >3 years.

¹² Hospital controls.

¹³ Neighborhood controls.

¹⁴ High quality studies.

¹⁵ Adjusted for smoking.

Crohn's Disease References

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Ulcerative Colitis

Papers were accessed from a PubMed literature review as noted (Williams 2017). Each paper was rated based on the parameters noted in the STROBE statement (von Elm 2007).

Overall 14 primary studies and one meta-analysis were identified which evaluated the effect of COCs on the later development of ulcerative colitis (Table 8). None of the primary studies has shown a statistically significant decrease in risk, while two showed a significant increase in risk for the development of ulcerative colitis with ever use of COCs (Boyko 1994, Parrello 1997). One meta-analysis examined ever use and failed to show a significant difference (Godet et al. 1995), while another meta-analysis examined current use and found a significantly increased relative risk of 1.28 (1.06–1.54). Overall these studies suggest that use of COCs conveys an increased risk of ulcerative colitis, especially current use.

Table 8 – Individual Studies of the Effects of COCs on the Development of Ulcerative Colitis

| Study | Study Design | OR Ever Use | RR Ever Use | OR Current Use | RR Current Use | OR Past Use | RR Past Use | Cases | Controls | Quality Score |
|-------------------------------------|--------------|---------------------|---------------------|--------------------|---------------------|---------------------|---------------------|-------|----------|---------------|
| Khalili et al. 2013 ¹ | Cohort | | 1.18 (0.92–1.52) | | 1.22 (0.74–2.07) | | 1.18 (0.91–1.52) | 392 | 116,983 | 93% |
| García Rodríguez et al. 2005 | Cohort | | | | 1.58 (0.71–3.52) | | 0.67 (0.32–1.39) | 222 | 10,000 | 88% |
| Logan and Kay 1989 | Cohort | | 1.3 (0.82–2.0) | | | | | 78 | 45,922 | 54% |
| Vessey et al. 1986 ² | Cohort | | | | 2.1 | | | 31 | 17,001 | 46% |
| Boyko et al 1994 | Case-control | | 1.7 (1.1–2.7) | | | | | 211 | 341 | 94% |
| Lashner et al. 1990 | Case-control | 0.86 (0.40–1.85) | | 0.7 (0.27–1.83) | | 1.14 (0.41–1.15) | | 46 | 46 | 81% |
| Sandler et al. 1992 ³ | Case-control | | 1.1 (0.65–1.85) | | | | | 89 | 217 | 81% |
| Persson et al. 1993 | Case-control | | 1.7 (0.8–3.3) | | | | | 145 | 305 | 81% |
| Halfvarson et al. 2006 ⁴ | Case-control | | | | 0.6 (0.1–2.5) | | | 125 | 125 | 75% |
| Ng et al. 2012 ⁵ | Case-control | 0.43 (0.11–1.66) | | | | | | 125 | 125 | 74% |
| Parrello et al. 1997 ⁶ | Case-control | 3.11 (1.54–6.3) | | | | | | 536 | 755 | 67% |
| Corrao et al. 1998 | Case-control | | | 1.6 (0.9–3.0) | | 1.3 (0.6–2.8) | | 594 | 594 | 67% |
| Calkins et al. 1986 ⁷ | Case-control | 0.62 (0.11–3.42) | | | | | | 35 | 32 | 42% |
| Calkins et al. 1986 ⁸ | Case-control | 0.57 (0.11–2.88) | | | | | | 35 | 38 | 42% |

¹ Hazard ratios (RR adjusted for time).

² Authors' calculation, adjusted for smoking.

³ Interaction with smoking notes, higher RR in smokers (2.49).

⁴ Monozygotic and dizygotic twins.

⁵ Twins studies.

⁶ Unclear how the calculation was done.

⁷ Hospital controls.

⁸ Neighborhood controls.

| Study | Study Design | OR Ever Use | RR Ever Use | OR Current Use | RR Current Use | OR Past Use | RR Past Use | Cases | Controls | Quality Score |
|----------------------------------|---------------|---------------------|---------------------|-------------------|----------------------|----------------|-----------------------|-------|----------|------------------|
| Vcev et al. 2015 | Case-control | 0.75 (0.30–1.88) | | | | | | 62 | 42 | 31% |
| Cornish et al. 2008 | Meta-analysis | | | | 1.28 (1.06–1.54) | | 1.07 (0.702–1.640) | 883 | 74,932 | 91% |
| Cornish et al. 2008 ⁹ | Meta-analysis | | | | 1.24 (0.999–1.54) | | | | | 91% |
| Godet et al. 1995 ¹⁰ | Meta-analysis | | 1.29 (0.94–1.77) | | | | | 851 | 49,875 | 82% |

⁹ High quality studies.

¹⁰ Adjusted for smoking.

Ulcerative Colitis References

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Systemic Lupus Erythematosus

Papers were accessed from a PubMed literature review as noted (Williams 2017). Each paper was rated based on the parameters noted in the STROBE statement (von Elm 2007).

There have been seven studies published evaluating the effect of hormonal contraceptives on susceptibility to systemic lupus erythematosus (Table 9). A significantly increased risk for development of systemic lupus erythematosus with use of COCs was shown for ever use in two studies (Costenbader 2007, Sanchez-Guerrero 1997), for current use in one study (Bernier 2009) and for past use in one study (Costenbader 2007). None of the studies showed a decreased risk. While no meta-analyses of these studies have been performed, the uniformity of the results implicate COCs as an important risk factor for the subsequent development of systemic lupus erythematosus.

Table 9 – Individual Studies of the Effects of COCs on the Development of Systemic Lupus Erythematosus

| Study | Study Design | OR Ever Use | RR Ever Use | OR Current Use | RR Current Use | OR Past Use | RR Past Use | Cases | Controls | Quality Score |
|---|--------------|-------------------|---------------------|-------------------|---------------------|------------------|---------------------|-------|----------|---------------|
| Costenbader et al. 2007 ¹ | Cohort | | 1.5 (1.1–2.1) | | | | 1.7 (1.2-2.3) | 262 | 238,046 | 96% |
| Costenbader et al. 2007 ² | Cohort | | 1.6 (1.1-2.2) | | | | 1.6 (1.1-2.2) | 164 | 102,882 | 96% |
| Costenbader et al. 2007 ³ | Cohort | | 2.3 (1.0-5.0) | | | | 2.3 (1.1-5.2) | 98 | 107,854 | 96% |
| Bernier et al. 2009 | Cohort | | 1.19 (0.98-1.45) | | 1.54 (1.15-2.07) | | 1.06 (0.85-1.33) | 786 | 7817 | 96% |
| Bernier et al. 2009 ⁴ | Cohort | | | | 2.52 (1.14-5.57) | | | 786 | 7817 | 96% |
| Bernier et al. 2009 ⁵ | Cohort | | | | 1.45 (1.06-1.99) | | | 786 | 7817 | 96% |
| Sanchez-Guerrero et al. 1997 | Cohort | | 1.4 (0.9-2.1) | | | | | 99 | 121,546 | 88% |
| Sanchez-Guerrero et al. 1997 ⁶ | Cohort | | 1.9 (1.1-3.3) | | | | | 58 | 121,587 | 88% |
| Cooper et al. 2002 | Case-control | | | 1.5 (0.8–2.7) | | 1.3 (0.8–2.0) | | 240 | 321 | 92% |
| Strom et al. 1994 | Case-control | 0.8 (0.5-1.4) | | | | | | 195 | 143 | 73% |
| Zonana-Nacach et al. 2002 ⁷ | Case-control | 2.1 (1.18-3.6) | | | | | | 130 | 130 | 61% |
| Grimes et al. 1985 | Case-control | | | 0.5 (0.11-2.3) | | | | 109 | 109 | 58% |

¹ Pooled RR from the Nurses' Health Study (NHS) and NHS II.

² RR from the NHS (data collection through 1976).

³ RR from NHS II (data collection through 1989).

⁴ RR for short term use (starting COCs within ≤3 months).

⁵ RR for long term use (starting COCs over 3 months previously with current use ongoing).

⁶ Using most stringent definition of systemic lupus erythematosus.

⁷ Paper written in Spanish. OR is for use of oral contraceptives for more than one year.

Systemic Lupus Erythematosus References

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Risk of Depression, Mood Disorders, and Suicide

The effects of contraceptive steroid hormones on depression, mood disorders, and suicide have been investigated (Table 10). The largest study of incident depression and use of anti-depressant medication (Skovlund 2016) indicates significantly increased risks for both COCs and POCs for both outcomes. The same group studied for suicide attempts and suicides (Skovlund 2018). Elevated risks were seen, and this was the case for both COCs and POCs. The recent NCHA study (Gregory 2018) showed a similar trend. One study (Keyes 2013) showed a lower risk of depression, but was not measuring clinically diagnosed depression, but rather the presence of depressive symptoms within 7 days prior to the survey. They also found a lower rate of suicide attempts among COC users. Similar findings were seen in 2 studies that also used a questionnaire looking at current COC or POC use (Toffol 2011, Toffol 2012). An analysis of the development of mood disorders found a higher incidence with POCs but a lower incidence with COCs (Svendal 2012). A study of post-partum depression as a reported adverse drug reaction showed higher rates for levonogestrel, etonogestrel and sertraline & drospirenone (Horibe 2018). A study of post-partum DMPA versus copper IUD use showed significant increases in depression scores and major depressive episodes with DMPA (Singata-Madliki, 2016). A retrospective cohort study showed increased risk for antidepressant use in patients who used ethinyl estradiol/etonogestrel (ring), and decreased risk of depression diagnosis with norethindrone-only pills or the levonorgestrel intrauterine system. A small retrospective chart review of the effect of immediate post-partum DMPA did not show significant effects on post-partum depression (Tsai 2009). All the papers, which have broken out the age groups of users, show maximum increased risk for depression, suicide risk, and suicide within 3 months of beginning to use the drugs and tapering off after 6 months, partly due to attenuation of symptoms, partly due to discontinuation due to adverse effects. These risks need to be adequately conveyed in prescribing information and patient-related materials.

However, little attention has been paid to the effects of blocking the important actions of estradiol and progesterone with progestins during the time of active brain remodeling. Estradiol and progesterone in normal sequence are essential for brain remodeling from ages 15–19 years particularly for myelination, dendritic pruning and establishment of new synaptic connections (Del Rio 2018). Suppressing these with synthetic progestins can have far-reaching, untoward effects. See Griksiene below in Table 10 as well as Del Rio (Del Rio 2018).

Table 10 – Studies of Chemical Contraceptives and Depression, Mood Disorders and Suicides

| | Study Design | OR Ever Use | RR Ever Use | OR Current Use | RR Current Use | Cases | Controls/Cohort Size |
|-------------------------------|---|----------------------------|------------------------------------|--|-------------------|---------|-------------------------|
| Skovlund 2016 incl /Worley | Prospective Cohort Incident Depression – COCs | | 1.1 ⁹⁵ (1.08-1.14) | | | | 1,061,997 |
| | Incident Depression – POCs | | 1.2 ⁹⁶ (1.04-1.31) | | | | |
| | First use of Antidepressants – COCs | | 1.23 ⁹⁷ (1.22-1.25) | | | | |
| | First use of Antidepressants – POCs | | 1.3 ⁹⁸ (1.27-1.40) | | | | |
| Skovlund 2018 incl /Worley | Prospective Cohort | | | | | | 475,802 |
| | Prospective Cohort Suicide attempts | | 1.97 ⁹⁹ (1.85-2.10) | | | | |
| | Suicides | | 3.08 ¹⁰⁰ (1.34-7.08) | | | | |
| Gregory 2018 | NCHA survey | | | | | 146,938 | 202,759 |
| | Ever Diagnosed with Depression | 1.558 (1.506- 1.612) | | | | | |
| | Academic performance affected by depression | 1.282 (1.245- 1.321) | | | | | |
| Keyes 2013 | COC reduced depression among women 25-34 years of age. ¹⁰¹ 4 waves of L-Hanes | | | -1.04 ¹⁰² (-1.73 - -0.35) | | 3224 | 1219 |
| | Suicide attempts | | | 0.38 (0.15-0.97) | | | |
| Toffol 2011 | Population/choice | | | -0.988 ¹⁰⁴ (-1.917 – -0.059) | | | 2,310 |

⁹⁵ First diagnosis of depression for combined oral contraceptive users.

⁹⁶ First diagnosis of depression for all progestin-only method users.

⁹⁷ First use of an antidepressant for combined oral contraceptive users.

⁹⁸ First use of an antidepressant for all progestin-only method users.

⁹⁹ Hazard ratio for suicide attempts; all hormonal contraceptives.

¹⁰⁰ Hazard ratio for suicides; all hormonal contraceptives.

¹⁰¹ “The presence of depressive symptoms during the past 7 days was assessed in all waves using the Center for Epidemiologic Studies Depression Scale (CES-D).”

¹⁰² β statistic shown.

¹⁰⁴ β statistic shown for the Beck Depression Inventory (BDI). None of the other parameters assessed was statistically significant (including any psychiatric diagnosis, alcohol dependence, major depressive episode or disorder, dysthymic disorder, or anxiety disorder).

| | Study Design | OR Ever Use | RR Ever Use | OR Current Use | RR Current Use | Cases | Controls/Cohort Size |
|-----------------------------|---|----------------|----------------|---|-------------------|--------------------|-------------------------|
| | Cross sectional 30-54 yrs. of age ¹⁰³ | | | | | | |
| Toffol 2012 | Population-based cross-sectional study ¹⁰⁵ | | | -0.42 (-1.79 - -0.04) ¹⁰⁶ | | | 8,586 |
| Svendal 2012 ¹⁰⁷ | Population-based cross-sectional study | | | | | 40 | 458 |
| | POC Use – mood disorder | | | 3.0 (1.1-7.8) | | | |
| | COC Use – mood disorder | | | 0.3 (0.1-0.9) | | | |
| Horibe 2018 | Retrospective ¹⁰⁸ | | | | | 253 | 6,157,897 |
| | Post-partum depression w/ levonorgestrel | | | 12.5 (8.7-18) | | | |
| | Post-partum depression w/ etonogestrel | | | 14.0 (8.5-22.8) | | | |
| | Post-partum depression w/ sertraline & drospirenone | | | 5.4 (2.7-10.9) | | | |
| Singata-Madliki 2016 | Single-blind randomized controlled trial of post-partum DMPA vs. copper IUD | | | 109 | | 111 ¹¹⁰ | 117 ¹¹¹ |

¹⁰³ “The associations between the current use of COCs and the LNG-IUS, and their duration versus mood symptoms [Beck Depression Inventory (BDI)], psychological well-being [(General Health Questionnaire-12 (GHQ-12))] and recent psychiatric diagnoses [(Composite International Diagnostic Interview (CIDI))] were examined among women who participated in the Finnish-population-based Health 2000 study.” “Overall, hormonal contraception was well tolerated with few significant effects on psychological well-being.”

¹⁰⁵ Data were collected in the context of the National FINRISK Study Survey, a cross-sectional population-based health survey carried out in Finland every 5 years since 1972. For the purpose of this study, data collected in the years 1997, 2002 and 2007 were analyzed for ages 25–54. OC vs. LNG. inconsistent questions between surveys, BDI, recall bias, etc. “Presence of somatic and psychological symptoms was assessed by asking the participants how often (often, sometimes, not at all) in the previous month they had had one or more out of 13 symptoms.” Also administered the Beck Depression Inventory-13. “A negative association between the current use of COCs and Beck Depression Inventory-13 (BDI-13) score was found. Some other negative associations, all characterized by a small effect size, were detected between current use of COCs and the BDI items feelings of dissatisfaction, feelings of uselessness, irritability, lost interest in people and lost appetite.”

¹⁰⁶ Results for the BDI-13 shown. Other parameters (including BDI-21, low mood last year, anhedonia last year, recent diagnosis of depression and recent other psychiatric diagnosis) did not reach statistical significance.

¹⁰⁷ Women in Australia 20-50 years of age. Evaluated for the occurrence of mood disorders, including major depressive disorder (MDD), minor depression, bipolar disorder, dysthymia, mood disorder due to a general medical condition and substance induced mood disorder.

¹⁰⁸ Data is from the FDA Adverse Event Reporting System (FAERS) database. Reporting Odds Ratios (ROR) are shown.

¹⁰⁹ Beck Depression Inventory (BDI-II) and the Edinburgh Postnatal Depression Scale (EPDS) evaluated. The one-month EPDS depression scores were statistically significantly higher in the DMPA arm compared with the IUD arm ($p=0.04$). Three-month BDI-II scores were significantly higher in the DMPA arm than in the IUD arm ($p=0.002$) and, according to the BDI-II but not the EPDS, more women in the DMPA arm had major depression at this time-point (8 vs 2; $p=0.05$).

¹¹⁰ 111 randomized to DMPA.

¹¹¹ 117 randomized to IUDs.

| | Study Design | OR Ever Use | RR Ever Use | OR Current Use | RR Current Use | Cases | Controls/Cohort Size |
|------------------------------|--|----------------|----------------|---|--|-----------------------|-------------------------|
| Kulkarni 2005 ¹¹² | Case-control pilot study COCs vs non-users | | | p=0.001 depression for all scales ¹¹³ | | 26 | 32 |
| Roberts 2017 | Retrospective cohort study ¹¹⁴ | | | With Dx of depression ¹¹⁵ | w/anti depressant use ¹¹⁶ | 31,506 ¹¹⁷ | 44,022 ¹¹⁸ |
| | Norethindrone-only pills | | | 0.56 (0.49-0.64) | 0.58 (0.52-0.64) | | |
| | Levonorgestrel intrauterine system | | | 0.65 (0.52-0.82) | 1.01 (0.87-1.18) | | |
| | Etonogestrel subdermal implant | | | 1.01 (0.83-1.22) | 1.22 (1.06-1.41) | | |
| | Ethinyl estradiol/ norgestimate (pill) | | | 0.89 (0.70-1.14) | 1.02 (0.85-1.22) | | |
| | Ethinyl estradiol/norethindrone (pill) | | | 0.82 (0.59-1.12) | 0.88 (0.69-1.13) | | |
| | Ethinyl estradiol/etonogestrel (ring) | | | 1.09 (0.80-1.50) | 1.45 (1.16-1.80) | | |
| Tsai 2009 | Retrospective chart review ¹¹⁹ | DMPA | Controls | | | 55 | 192 |
| | Mean EPDS scores at 6 weeks postpartum | 5.02 | 6.17 | | | | |
| Griksiene 2011 | Case-control study ¹²⁰ | ¹²¹ | | | | 23 ¹²² | 20 ¹²³ |

¹¹² Assessment tools included three depression rating scales: Beck Depression Inventory (BDI), Hamilton Rating Scale for Depression (HAM-D) and Montgomery-Asberg Depression Rating Scale (MADRS); also used the Global Assessment of Functioning (GAF) Scale.

¹¹³ ANOVA of GAF, BDI, HAM-D & MADRS scales all significantly different.

¹¹⁴ Post-partum depression with hormonal contraception.

¹¹⁵ Adjusted hazard ratios shown.

¹¹⁶ Adjusted hazard ratios shown.

¹¹⁷ Number on hormonal contraceptives.

¹¹⁸ Number not on hormonal contraceptives.

¹¹⁹ Depot medroxyprogesterone in the immediate post-partum period and depression. Evaluated the Edinburgh Postnatal Depression Scale (EPDS).

¹²⁰ Verbal fluency and mental rotation (spatial perception) are affected by progestins w/androgenic or antiandrogenic properties.

¹²¹ Naturally cycling women performed better on verbal fluency task as compared to OC users. Subjects who used the third generation (androgenic) COCs generated significantly fewer words as compared to new generation (anti-androgenic) OC users and non-users. The third generation OC users demonstrated significantly longer RT in MRT task as compared to non-users. The MRT, verbal fluency and mood parameters did not depend on the phase of menstrual cycle.

¹²² Women on hormonal contraception.

¹²³ Control women not on hormonal contraception.

Depression, Mood Disorders, and Suicide References

Del Rio JP, Allende MI, Molina N, Serrano FG, Molina S, and Vigil P. Steroid Hormones and their Action in Women's Brains: The Importance of Hormonal Balance. *Frontiers in Public Health* 2018; May(6) art. 141:1–15.

Gregory Sean T, Hall K, Quast T, Gatto A, Bleck J, Storch EA, and DeBate R. Hormonal contraception, depression and Academic Performance among females attending college in the United States. *Psychiatry Research* 2018; 270:111–116.

Keyes Katherine T, Cheslack-Postava K, Westhoff C, Heim CM, Haloosim M, Walsh K, and Koenen K. Association of Hormonal Contraception Use with Reduced levels of Depressive Symptoms: A National Study of Sexually Active women in the United States. *Am J. Epidemiol* 2013; 178(9):1378–1388.

Skovlund CW, Mørch LS, Kessling LV, and Lidegaard O. Association of Hormonal Contraception with Depression. *JAMA Psychiatry* 2016; 73(11):1154–1162.

Skovlund CW, Mørch LS, Kessling LV, Lange T, and Lidegaard, O. Association of Hormonal Contraception with Suicide Attempts and Suicides. *Am. J Psychiatry* 2018; 175(4):336–342.

Svendal G, Berk M, Pasco JA, and Jacka FN. The use of hormonal contraceptive agents and mood disorders in women. *J Affective Disorders* 2012; 140:92–96.

Toffol E, Heiknheimo, Koponene P, Luoto R, and Partonen T. Hormonal contraception and mental health: results of a population based study. *Human Reproduction* 2011; 26(11):3085–3093.

Worly Brett L, Gur TL, and Schaffir J. The relationship between progestin hormonal contraception and depression: a systematic review. *Contraception* 2018; 97:478–489.

Young EA, Kornstein SG, Harvey AT, Wisniewski SR, Barkin J, Fava M, Trivedi MH, and Rush AJ. Influences of Hormone-Based Contraception in Depressive symptoms in Premenopausal Women with Major Depression. *Psychoneuroendocrinology* 2007; 32(7):843–853.

Multiple Sclerosis

Papers were accessed from a PubMed literature review as noted (Williams 2017). Each paper was rated based on the parameters noted in the STROBE statement (von Elm 2007).

A total of 6 studies (3 cohort studies and 3 case-control studies) were identified which evaluated the impact of COCs on the subsequent development of multiple sclerosis (Table 11). Two studies showed a significantly increased risk for the development of multiple sclerosis with ever use of COCs (Hellwig 2016, Kotzamani 2012) with a similarly increased risk noted in one study for current use or past use (Hellwig 2016). Overall these studies suggest that use of COCs may convey an increased risk for the subsequent development of multiple sclerosis.

Table 11 – Individual Studies of the Effects of COCs on the Development of Multiple Sclerosis

| Study | Study Design | OR Ever Use | RR Ever Use | OR Current Use | RR Current Use | OR Past Use | RR Past Use | Cases | Controls | Quality Score |
|--------------------------------------|--------------|---------------------|------------------|-------------------|------------------|---------------------|------------------|-------|----------|---------------|
| Hernán et al. 2000 ¹²⁴ | Cohort | | 1.1 (0.9-1.5) | | 1 (0.6-1.6) | | 1.2 (0.9-1.5) | 313 | 237,318 | 90% |
| Thorogood et al. 1998 ¹²⁵ | Cohort | | | | 1.2 (0.7-2.0) | | 1.3 (0.9-2.0) | 114 | 46,000 | 75% |
| Villard-Mackintosh et al. 1993 | Cohort | | 0.8 (0.5-1.4) | | | | | 63 | 16,969 | 65% |
| Hellwig et al. 2016 | Case-control | 1.51 (1.12-2.03) | | 1.47 1.05-2.05 | | 1.55 (1.20-2.00) | | 400 | 3804 | 92% |
| Kotzamani et al. 2012 | Case-control | 1.6 (1.1-2.4) | | | | | | 254 | 314 | 81% |
| Alonso et al. 2005 ¹²⁶ | Case-control | 0.6 (0.4-1.0) | | 0.5 (0.3-1.2) | | 0.6 (0.4-1.0) | | 106 | 1001 | 77% |

¹²⁴ NHS I and II cohorts.

¹²⁵ Funded by drug companies that make HCs.

¹²⁶ OC use over the 3 years prior to the index date. Limited to women ≤50 years of age.

Multiple Sclerosis References

Alonso A, Jick SS, Olek MJ, Ascherio A, Jick H, and Hernán MA. Recent use of oral contraceptives and the risk of multiple sclerosis. *Archives of Neurology* 2005; 62:1362–1365.

Hellwig K, Chen LH, Stanczyk FZ, and Langer-Gould AM. Oral Contraceptives and Multiple Sclerosis/Clinically Isolated Syndrome Susceptibility. *PLoS One* 2016; 11:e0149094. Doi:10.1371/journal.pone.0149094.

Hernán MA, Hohol MJ, Olek MJ, Spiegelman D, and Ascherio A. Oral contraceptives and the incidence of multiple sclerosis. *Neurology* 2000; 55:848–854.

Kotzamani D, Panou T, Mastorodemos V, Tzagournissakis M, Nikolakaki H, Spanaki C, and Plaitakis A. Rising incidence of multiple sclerosis in females associated with urbanization. *Neurology* 2012; 78:1728–1735.

Thorogood M, and Hannaford PC. The influence of oral contraceptives on the risk of multiple sclerosis. *British Journal of Obstetrics and Gynaecology* 1998; 105:1296–1299.

Villard-Mackintosh L, and Vessey MP. Oral contraceptives and reproductive factors in multiple sclerosis incidence. *Contraception* 1993; 47:161–168.

Interstitial Cystitis

A case-control study (Konkle 2012) showed significantly higher use of birth control pills in cases versus controls: 88% versus 82%; $P = 0.019$. Another case-control study showed that use of COCs markedly increased the risk of the disease whether past (OR 4.6, 95% CI 1.74-12.1) or current use (OR 6.9, 95% CI 2.1–22.1). Interstitial cystitis was associated with vulvodynia and sexual dysfunction in a high number of cases (Gardella 2011). Another study showed that use of COCs in patients with interstitial cystitis was associated with a decrease in quality of life (El Khoudary 2009). One meta-analysis (Champaneria 2015) showed that ever use of COCs significantly increased the risk of interstitial cystitis (OR 2.31, 95% CI 1.03–5.16).

Overall, use of COCs appears to be associated with an increased risk for the development of interstitial cystitis.

Interstitial Cystitis References

Champaneria R, D'Andrea RM, and Latthe PM. Hormonal contraception and pelvic floor dysfunction: a systematic review. *Int Urogynecol J* 2016; 27:709–722.

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Osteoporotic Bone Fractures

Prescribing information for POCs typically includes a warning regarding the development of osteoporosis. However, the more relevant outcome is fracture risk. Therefore, articles were sought that looked at the effect of COCs and POCs on fracture risk. Data were initially derived from a systematic review of the evidence from observational studies of hormonal contraceptive use for contraception and the risk of fracture in women by Lopez (Lopez 2015). They noted that in 2004, the US Food and Drug Administration added a warning to depot medroxyprogesterone acetate (DMPA) labeling about the potential loss of BMD (FDA 2004), which might limit long-term use. A systematic review of progestin-only methods found an association between DMPA use and loss of bone mineral density (Curtis 2006). Lopez identified 559 records, 524 of which did not meet their inclusion criteria. Thirty-five full-text reports remained, 11 of which were excluded. Of the remaining 24, 10 were secondary articles. That left 14 articles: the 14 studies examined oral contraceptives (N = 12), DMPA (N = 4) and the hormonal IUD (N = 1). Similar search terms to Lopez were used for papers published since 2015 and 2 additional papers were retrieved. The resulting studies are shown in Table 12.

COCs: Three early studies (Cooper 1993, Tuppurainen 1993, Vessey 1998) showed an increase risk of fracture with use of COCs. These studies predominately evaluated pre-menopausal fracture risk. Others that evaluated wrist fracture linked to falling had few cases but showed a trend to decreased risk (O'Neill 1996). One study that evaluated post-menopausal fracture risk based on prior oral contraceptive use (Barad 2005) also found an increased fracture risk. Another study looking at hip fracture risk in elderly women (Michaëlsson 1999) showed a decreased risk but is compromised in that "The exposure time for oral contraceptives may thus maximally have spanned 5 years..." Two studies by Vestergaard (Vestergaard 2006 and Vestergaard 2008) looked at any fracture with OC use and did not show a significant effect when multivariate analyses were performed. However, these studies only looked at use within the past 5 years and did not take into account remote use or cumulative lifetime use. A small cross-sectional study in southern Tasmania (Wei 2011) was stratified by duration of use and showed a reduction in vertebral deformities for 5-10 years of use, but no effect for shorter or longer duration of use and no effect on number of vertebral deformities. A large case-control study which evaluated incident fracture risk with varying numbers of COC prescriptions showed an increased risk for 10+ prescriptions with current use (Meier 2010). A similar study failed to confirm this for most prescription numbers (Kyvernitakis 2016) but this study had fewer subjects reducing its power. A case-control study (Memon 2011) nested in an earlier cohort study (Cooper 1993) failed to show an effect.

Overall the weight of evidence for use of COCs suggests an increased risk of bone fracture with protracted use. The study by Barad (2005) appears to have the largest number of subjects, was a cohort study, and was the only study that evaluated post-menopausal fracture risk with prior use of COCs.

In contrast, virtually all the studies evaluating POCs show an elevated risk (Lanza 2013, Vestergaard 2008b, Meier 2010, Kyvernitakis 2016). This risk appears to increase with duration of use.

Table 12 – Individual Studies of the Effects of Contraceptives on the Development of Osteoporotic Fractures

| Study | Study Design | Intervention | OR | RR | Cases | Controls or Cohort Size | Outcome |
|----------------------------|----------------------------|---------------------|----|---------------------|--------|-------------------------|-----------------------------------|
| Cooper 1993 ¹²⁷ | Cohort | COCs | | 1.20 (1.08-1.34) | 1365 | 46,000 | All fractures |
| Vessey 1998 ¹²⁸ | Cohort | COCs | | 1.5 (1.1-2.1) | 1308 | 17,032 | First fracture: radius or ulna |
| Vessey 1998 ¹²⁹ | Cohort | COCs | | 1.2 (1.1-1.4) | | | First fracture: all sites |
| Vessey 1998 ¹³⁰ | Cohort | COCs | | 2.5 (1.5-4.0) | | | First fracture: radius or ulna |
| Vessey 1998 ¹³¹ | Cohort | COCs | | 1.3 (1.1-1.5) | | | First fracture: all sites |
| Vessey 1998 ¹³² | Cohort | COCs | | 5.7 (p=0.017) | | | First fracture: radius or ulna |
| Vessey 1998 ¹³³ | Cohort | COCs | | 11.2 (p<0.001) | | | First fracture: all sites |
| Barad 2005 ¹³⁴ | Cohort | OCs ¹³⁵ | | 1.07 (1.01–1.15) | 4,674 | 80,947 | First fracture |
| Barad 2005 ¹³⁶ | Cohort | OCs | | 1.15 (1.04-1.27) | 4,674 | 80,947 | First fracture |
| Barad 2005 ¹³⁷ | Cohort | OCs | | 1.09 (0.97–1.23) | 4,674 | 80,947 | First fracture |
| Lanza 2013 ¹³⁸ | Retrospective cohort study | DMPA ¹³⁹ | | 1.41 (1.35–1.47) | 11,822 | 312,395 | Incident fractures |

¹²⁷ From the Royal College of General Practitioners (RCGP) Oral Contraception Study.

¹²⁸ OC use > 97 months vs no use. Recruited age 25 to 39 years; followed to 45 years.

¹²⁹ OC use > 97 months vs no use. Recruited age 25 to 39 years; followed to 45 years.

¹³⁰ Interval since use: 73 to 96 months vs no use (radius or ulna). Recruited age 25 to 39 years; followed to 45 years.

¹³¹ < 12 months vs no use (all fractures). Recruited age 25 to 39 years; followed to 45 years.

¹³² χ^2 trend.

¹³³ χ^2 trend.

¹³⁴ Recruited age 50 to 74 years; OC use: any vs none.

¹³⁵ The patients were asked about oral contraceptive use, which likely was predominantly COCs but was not broken down with regard to COCs or POCs.

¹³⁶ Among women without any postmenopausal hormone treatment, past OC use for 5 years or less.

¹³⁷ Among women without any postmenopausal hormone treatment, past OC use for more than 5 years.

¹³⁸ They note that, “Although DMPA users experienced more fractures than nonusers, this association may be the result of confounding by a pre-existing higher risk for fractures in women who chose DMPA for contraception.” However, this is based on analysis of relatively few fractures prior to DMPA use.

¹³⁹ Depot medroxyprogesterone acetate = DMPA.

| Study | Study Design | Intervention | OR | RR | | Cases | Controls or Cohort Size | Outcome |
|----------------------------------|----------------------------|--------------------|---------------------|---------------------|---------------------|--------|-------------------------|--|
| | Past use ¹⁴⁰ | DMPA | | 1.32 (1.24–1.41) | | | | Incident fractures |
| | Recent use ¹⁴¹ | DMPA | | 1.41 (1.31–1.50) | | | | Incident fractures |
| | Current use ¹⁴² | DMPA | | 1.51 (1.41–1.61) | | | | Incident fractures |
| Tuppurainen 1993 ¹⁴³ | Case-control | OCs | 1.21 (0.93–1.57) | | | 629 | 13,100 | All fractures |
| Tuppurainen 1993 ¹⁴⁴ | Case-control | OCs | 1.35 (0.88–2.05) | | | 210 | 13,100 | Wrist fractures |
| O’Neill 1996 | Case-control | OCs | 0.3 (0.1–0.9) | | | 62 | 116 | Distal forearm fractures only Population controls |
| O’Neill 1996 | Case-control | OCs | 0.7 (0.2–2.4) | | | 62 | 50 | Distal forearm fractures only Fall controls |
| Michaëlsson 1999 ¹⁴⁵ | Case-control | Any ¹⁴⁶ | 0.75 (0.59–0.96) | | | 1327 | 3312 | Hip fractures |
| Vestergaard 2006 ¹⁴⁷ | Case-control | OCs | <0.3 DDD/day | 0.3–0.99 DDD/day | 1+ DDD/day | 64,548 | 193,641 | Any fracture in the year 2000 |
| | <25 years ¹⁴⁸ | OCs | 0.97 (0.91–1.03) | 0.96 (0.92–1.01) | 0.92 (0.86–0.98) | | | Any fracture in the year 2000 |
| | 25–49 years | OCs | 0.91 (0.82–1.00) | 0.90 (0.77–1.05) | 0.87 (0.64–1.18) | | | Any fracture in the year 2000 |
| | 50+ years | OCs | 0.92 (0.77–1.10) | 0.69 (0.45–1.05) | 0.62 (0.27–1.41) | | | Any fracture in the year 2000 |
| Vestergaard 2008a ¹⁴⁹ | Case-control | OCs | <0.3 DDD/day | 0.3–0.99 DDD/day | 1+ DDD/day | 64,548 | 193,641 | Any fracture in the year 2000 |
| | <15 | OCs | 1.02 (0.75–1.37) | 1.17 (1.01–1.37) | 0.97 (0.85–1.11) | | | Any fracture in the year 2000 |

¹⁴⁰ Active DMPA use based on the interleaving of active 90-day exposures generated by each injection.

¹⁴¹ Recent exposure is 640 or fewer days after the last active exposure.

¹⁴² Past exposure begins after “recent” exposure (641 or more days after the last active exposure).

¹⁴³ Oral contraceptive use for 6+ years.

¹⁴⁴ Oral contraceptive use for 6+ years.

¹⁴⁵ No significant correlation was seen with duration of use, time since last use or time between last use and menopause.

¹⁴⁶ Any type of chemical contraceptive was evaluated, not separated as COCs or POCs.

¹⁴⁷ “The exposure time for oral contraceptives may thus maximally have spanned 5 years (from January 1, 1996, to December 31, 2000).” This and the other Vestergaard study are not useful as they do not take into account remote use or cumulative lifetime use. ORs shown.

¹⁴⁸ Defined daily dosages = DDD.

¹⁴⁹ Similar to Vestergaard 2006; only looked at use within the past 5 years. A younger group examined here. ORs shown.

| Study | Study Design | Intervention | OR | RR | | Cases | Controls or Cohort Size | Outcome |
|----------------------------------|------------------|--------------|---------------------|---------------------|---------------------|--------|-------------------------|---|
| | 15.1-17 | OCs | 1.22 (1.02–1.47) | 1.14 (1.00–1.30) | 1.04 (0.90–1.19) | | | Any fracture in the year 2000 |
| | 17.1-19 | OCs | 0.97 (0.87–1.09) | 0.93 (0.84–1.02) | 1.02 (0.89–1.18) | | | Any fracture in the year 2000 |
| | >19 | OCs | 0.99 (0.93–1.05) | 1.00 (0.93–1.08) | 0.88 (0.78–0.99) | | | Any fracture in the year 2000 |
| Vestergaard 2008b ¹⁵⁰ | Case-control | DMPA | 1.44 (1.01–2.06) | | | 64,548 | 193,641 | Any fracture in the year 2000 DMPA use |
| Wei 2011 ¹⁵¹ | Cross-sectional | | <5 years of use | 5-10 years of use | >10 years of use | | 491 | |
| | | OCs | 0.85 (0.45–1.58) | 0.45 (0.21–0.93) | 0.75 (0.36–1.54) | | | Presence of vertebral deformity |
| | | OCs | 0.96 (0.62–1.48) | 0.63 (0.37–1.07) | 0.94 (0.56–1.56) | | | Number of vertebral deformities |
| Meier 2010 ¹⁵² | Case-control | | Current Use | Past Use | | 17,527 | 70,130 | Incident fracture |
| | 1-2 DMPA Scripts | DMPA | 1.18 (0.93–1.49) | 1.17 (1.07–1.29) | | | | Incident fracture |
| | 3-9 DMPA scripts | DMPA | 1.36 (1.15–1.60) | 1.23 (1.11–1.36) | | | | Incident fracture |
| | 10+ DMPA scripts | DMPA | 1.54 (1.33–1.78) | 1.30 (1.09–1.55) | | | | Incident fracture |
| | 1-2 COC Scripts | COCs | 1.01 (0.87–1.18) | 1.00 (0.95–1.07) | | | | Incident fracture |
| | 3-9 COC scripts | COCs | 1.01 (0.94–1.09) | 0.99 (0.94–1.04) | | | | Incident fracture |
| | 10+ COC scripts | COCs | 1.09 (1.03–1.16) | 1.03 (0.97–1.10) | | | | Incident fracture |
| Memon 2011 ¹⁵³ | Case-control | COCs | 1.05 (0.86–1.29) | | | 651 | 1302 | Any fracture |
| Kyvernitakis 2016 ¹⁵⁴ | Case-control | | OR Current Use | OR Past Use | | 4189 | 4189 | First-time fracture diagnosis |
| | 1-2 DMPA scripts | DMPA | 0.97 (0.51–1.86) | 0.96 (0.73–1.26) | | | | |
| | 3-9 DMPA scripts | DMPA | 2.41 (1.42–4.08) | 1.14 (0.86–1.51) | | | | |
| | 10+ DMPA scripts | DMPA | 1.46 (0.96–2.23) | 1.55 (1.07–2.27) | | | | |

¹⁵⁰ Similar to Vestergaard 2006; only looked at use within the past 5 years. DMPA examined here. ORs shown.

¹⁵¹ Small cross-sectional study. ORs shown.

¹⁵² Females aged 20–44 years with an incident fracture diagnosis between 1995 and 2008.

¹⁵³ Nested case-control study of the Cooper study from the Royal College of General Practitioners (RCGP) Oral Contraception Study. Last OC use > 10 years vs never.

¹⁵⁴ Women between 20 and 44 years of age with a first-time fracture diagnosis, matched with random controls using the Disease Analyzer database.

| Study | Study Design | Intervention | OR | RR | | Cases | Controls or Cohort Size | Outcome |
|-------|-----------------|--------------|---------------------|---------------------|--|-------|-------------------------|---------|
| | 1-2 COC scripts | COCs | 0.98 (0.73–1.31) | 0.90 (0.77–1.05) | | | | |
| | 3-9 COC scripts | COCs | 1.39 (1.12–1.73) | 0.90 (0.78–1.03) | | | | |
| | 10+ COC scripts | COCs | 1.07 (0.88–1.30) | 1.04 (0.90–1.21) | | | | |

Osteoporotic Fracture References

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Impact of Contraceptives on Body Mass

Weight gain is a common complaint among contraceptive users but whether use of contraceptives is causally related remains undefined. Progestin-only contraceptives are most commonly associated with weight gain complaints and discontinuation. A recent Cochrane review (Gallo et al. 2014) examined the effect of combined oral contraceptives on weight gain and concluded existing data does not support a causal relationship. A second review of progestin-only contraceptives on weight gain (Lopez et al. 2016) found most studies of low to moderate quality but did conclude weight gain of up to 2kg (4.4 lbs) within the first year of use with continued increases thereafter. The authors advised appropriate counselling on expected weight changes to minimize discontinuation due to perceived weight gain.

The attached table (Table 13) summarizes studies of 1 year or longer that examined weight and body mass changes in contraceptive users in comparison to non-hormonal contraceptives or no method. Several additional studies compare various contraceptives for their effect on weight or body composition, but these do not directly address our focus.

The strongest data appear to be the deleterious effects of levonorgestrel-releasing IUDs on percent lean and fat body mass. Total body weight change does not appear different between groups and several large studies have shown no significant differences. However, a significant increase in % fat mass with a corresponding decrease in % lean body mass was observed in both studies where these were measured. A similar effect was seen from oral desogestrel in a single study.

Thus, while limited to date, data suggest that use of progestin-only contraceptives may have deleterious effects on % fat and % lean body mass with no significant overall effect on total body weight.

A review of current Mirena labeling makes no mention of changes in lean or fat body mass composition.

Retrospective, but not more recent, prospective studies also show DMPA use is associated with significant gains in weight. The data appear too mixed to draw firm conclusions.

Table 13 – Effect of Chemical Contraceptives on Weight Gain

| Study | Design | Comparison | N | Time | Weight change (Kg) | Fat mass change | Lean mass change | Comments |
|--|------------|-----------------------------|------|------|--------------------|-------------------|-----------------------|---|
| Pantoja 2010 | Retrospec. | DMPA 150 vs CuIUC | 758 | 1yr | 1.76 vs-0.42* | | | Largest differences noted in normal and overweight BMI subgroups, minimal differences in obese BMI subgroup |
| | | | | 2yr | 3.1 vs 0.4* | | | |
| | | | | 3yr | 3.9 vs 0.8* | | | |
| Modesto 2015 | Retrospec. | DMPA150 vs CuIUC | 1277 | 1yr | 1.3 vs 0.2* | | | Adjusted for years of school & # children. 20% loss @4yrs 84% @ 10yr. |
| | | | | 4yr | 3.5 vs 1.9* | | | |
| | | | | 10yr | 6.6 vs 4.9* | | | |
| Taneepanichskul 1998 | Retrospec. | DMPA 150 vs CuIUC | 100 | 10yr | 10.9 vs 11.2 | | | Included women 37-50 years (no younger women) |
| Vickery 2013 | Prospec. | DMPA 150 vs CuIUC | 167 | 1yr | 2.2 vs 0.16 | | | CHOICE study subgroup |
| Dal'Ava 2014 | Prospec. | DMPA 150 vs CuIUC | 110 | 1yr | 1.9vs 1.1 | 1.6 vs -0.9 (Kg) | 0.3 vs 1.2 (kg) | Paired by age (+/-2yr) & weight (+/-2kg) |
| Dos Santos 2014 | Prospec. | DMPA 150 vs CuIUC | 71 | 1yr | 1.4 vs 0.3 | 1.57 vs 0.52 (kg) | (0.31) vs (0.26) (kg) | Matched by age & BMI (-)= negative value |
| Studies comparing LNG IUC to non-hormonal contraceptive | | | | | | | | |
| Study | Design | Comparison | N | Time | Weight change (Kg) | Total body fat | Lean body mass | |
| Dal'Ava 2012 | Prospec. | LNG-IUC vs non-hormonal IUC | 76 | 1yr | 2.9 vs 1.4 | 2.5% vs -1.3%* | (1.4%) vs 1.0%* | Paired by age & BMI |
| Napolitano 2015 | Prospec. | LNG IUC vs no method | 60 | 1yr | 0.6 vs (0.2) | 1.1% vs (0.5%)* | (1.1%) vs 0.5* | |
| Vickery 2013 | Prospec. | LNG-IUC vs Cu IUC | 230 | 1yr | 1.03 vs 0.16 | nd | nd | |

| | | | | | | | | |
|--------------|------------|---------------------|------|------|------------|----|----|--|
| Modesto 2015 | Retrospec. | LNG-IUC vs CuIUC | 1204 | 1yr | 0.7 vs 0.2 | nd | nd | |
| | | | | 4yr | 2.7 vs 1.9 | | | |
| | | | | 10yr | 4.0 vs 4.9 | | | |
| | | | | | | | | |

Studies comparing progestin-only COCs to non-hormonal

| Study | Design | Comparison | N | Time | Weight change (Kg) | Total body fat | Lean body mass | |
|-----------------|----------|------------------------------------|----|------|--------------------|-------------------|-----------------|--|
| Napolitano 2015 | Prospec. | Desogestrel 75ug vs no hormonal | 68 | 1yr | 0.3 vs -0.2 | 1.1% vs -0.5%* | (2.8%) vs 0.5%* | |
| | | | | | | | | |
| | | | | | | | | |

Studies comparing combined COCs to non-hormonal

None found-

Abstract from 2014 Cochrane review of combined oral contraceptives on weight gain:

"We found 49 trials that met our inclusion criteria. The trials included 85 weight change comparisons for 52 distinct contraceptive pairs (or placebos). *The four trials with a placebo or no intervention group did not find evidence supporting a causal association* between combination oral contraceptives or a combination skin patch and weight change. Most comparisons of different combination contraceptives showed no substantial difference in weight. In addition, discontinuation of combination contraceptives because of weight change did not differ between groups where this was studied.

Gallo MF, Lopez LM, Grimes DA, Carayon F, Schulz KF, Helmerhorst FM. Combination contraceptives: effects on weight. Cochrane Database of Systematic Reviews 2014, Issue 1. Art. No.: CD003987.

* Significant difference (p<0.05).

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Urogenital Effects of Contraceptives

In addition to cervical cancer and interstitial cystitis, noted above, there are other adverse urogenital effects of COCs that should be communicated to patients. These include bacteriuria (Zahran 1976; calculated OR 3.57), urinary tract infection (Engel 1979: 27–50% incidence), bladder trabeculation (Zahran 1976; calculated OR 11.7), recurrent vulvovaginal candidiasis (Spinillo 1995, Yusuf 2007; OR 2.08), vaginal dryness (Lee 2017), vulvar vestibulitis (Champaneria 2016: OR 2.1 95 % CI 1.26–3.49; also noted in Lee 2017), and Female Sexual Dysfunction (FSD) (Lee 2017). FSD appears related to OC-induced dyspareunia, reduced sexual desire and libido (Lee 2017). This risk is increased if COCs are used in adolescents and the duration of OC use is at least 2 years (Lee 2017), although some newer COCs containing drospirenone 3 mg plus EE 30 mg and gestodene 75 mg plus EE 20 mg appear to have a reduction in these risks (Lee 2017).

These urogenital risks, especially FSD where there is substantial literature, should be referenced in prescribing information and patient pamphlets.

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Venous Thromboembolism and Contraceptives

The current language on the black box warning of certain contraceptives regarding risk of cardiovascular events clearly misleads women about the real risks of these drugs. It says: WARNING: CIGARETTE SMOKING AND SERIOUS CARDIOVASCULAR EVENTS. A study (Gomer 2009) conducted among 300 women concluded “that most of them believe that certain risks are only associated with being over 35 years of age and/or smoking.” Instead, the label should clearly state that anyone taking the medications without good knowledge of the risk factors could experience a potentially life-threatening cardiovascular event and should discuss the risks with a medical provider.

The incidence of venous thromboembolism (VTE) for healthy women can significantly increase with the use of hormonal contraceptives, even women under 35 and not-smoking. In a 2012 article about birth control side effects, Dr. Rebecca Peck (Peck R 2012) reports that “Oral contraceptives are associated with a three to five times higher risk of VTE (Van Hylckama VA 2009).” Third and fourth generation combined hormonal contraceptives (CHC) have been found to put women at an even much higher risk, leading to major lawsuits against some manufacturers and changes in regulations in several countries. In his opinion published in Drug Safety, Dr. Lidegaard, the author of several studies on the subject, states: “Of 14 studies specifically assessing the risk in users of CHC with desogestrel or gestodene, 13 found a higher risk with use of these products when compared to the use of CHC with levonorgestrel” (Lidegaard 2014). Drospirenone, the progestin contained in Yaz and Jasmine, also increases the risk of VTE over levonorgestrel by a factor of 1.5 to 2.8. “The relative risk [of Drospirenone was 6.3 as compared with nonusers in both the large Dutch (Van Hylckama 2009) and Danish (Lidegaard 2011) study.” The author comments that “the studies demonstrating risk differences between CHC with different progestins are generally methodologically more transparent and more robust than those demonstrating no difference, especially concerning exclusion of women with predispositions for VTE.” Another large study published in 2015 (Vinogradova 2015) reviewed 10,552 cases of VTE reported between 2001 and 2013 in the UK and found similar elevated risks of VTE with these CHC: “Corresponding risks associated with current exposure to desogestrel (4.28, 3.66 to 5.01), gestodene (3.64, 3.00 to 4.43), drospirenone (4.12, 3.43 to 4.96), and cyproterone (4.27, 3.57 to 5.11) were significantly higher than those for second generation contraceptives levonorgestrel (2.38, 2.18 to 2.59).” Note that the odds ratios were “adjusted for smoking status, alcohol consumption, ethnic group, body mass index, comorbidities, and other contraceptive drugs.”

Most importantly, the risk levels are multiplied if women have other risk factors. For instance, women who have the genetic blood condition known as Factor V Leiden could have a risk as high as 18 per 10,000 woman-years. If these women stay on the product for 10 years, their risks could be 250 per 10,000 woman-years, or 2.5% as risks increase with aging (Lidegaard 2014).

Dr. Lidegaard concludes: “Therefore, women with known risk factors of VTE are advised to be reluctant to use CHC. The relative risk of VTE with different dispositions is as follows: previous thrombosis: > 50 (Le Moigne 2013), genetic abnormalities such as factor V Leiden mutation (heterozygous): 6, deficiency of protein C: 10, of protein S: 10, of antithrombin: 25, and of prothrombin 20210A: 3 (Phillippe 2014). Pregnancy with delivery on average: 8, adiposity: 2–3 and immobilization 2–5 depending on how long time you are immobilized. Family disposition (first-degree relatives with VTE before their 50th year) doubles the risk of VTE. Women with such dispositions are generally recommended to use progestin-only contraception, which does not increase the risk of VTE except perhaps for medroxyprogesterone depots. A genetic screening should until further also be restricted to women with a family disposition” (Lidegaard 2014).

In a 2018 systematic review (Keenan 2018) of the most evidenced-based articles from the 1960s to 2018 comparing users of COCs to nonusers, with a confirmed diagnosis of VTE, and including more than 17 million

woman-years of observation, women on HC increase their risk by 3- to 9-fold. However, the first year of use has the highest risk for clot formation, and if a woman is younger than 30, her risk is increased 13-fold in the first year. Obesity can increase the risk of being on hormonal contraception, about doubling the risk compared to a woman of normal weight on the pill. It is not considered cost-effective to check for thrombophilia, a genetic disposition to form blood clots, but for those with thrombophilia, the risk can be as high as 62-fold in the first year.

This systematic review of the literature concludes that 136–260 women die from VTE a year in the United States from hormonal contraception. Combined with the added risk of stroke and heart attack from the COCs, 300–400 women die each year in the United States simply due to their choice of using HC for family planning (Keenan 2018). To give some perspective, meningitis killed 45 people (of all ages) in 2017: most US States mandate meningitis vaccination for college and university students.

A summary of studies is shown in Table 14.

Table 14 – Relative Risk of Venous Thromboembolism in Current Users of Different Combined Hormonal Contraceptives as Compared with Nonusers Unless Otherwise Specified

| Study | Data Sampling Period | VTE (number) | CHCs with levonorgestrel RR (95% CI) | CHCs with desogestrel/gestodene RR (95% CI) | CHCs with drospirenone RR (95% CI) |
|------------------|----------------------|--------------|---|--|---------------------------------------|
| Blomenkamp 1995 | 1988 - 1992 | 126 | 3.8 (1.7 - 8.4) | 8.7 (3.9 - 19.3) | - |
| WHO 1995a, 1995b | 1989 - 1993 | 433 | 3.6 (2.5 - 5.1) | 7.4 (4.2 - 12.9) | - |
| Jick 1995 | 1991 - 1994 | 80 | 1 (reference) | 1.8 (1.0 - 3.2) | - |
| Spitzer 1996 | 1991 - 1995 | 471 | 3.7 (2.2 - 6.2) | 6.7 (3.4 - 13.0) | - |
| Lewis 1999 | 1993 - 1995 | 502 | 2.9 (1.9 - 4.2) | 2.3 (1.5 - 3.5) | - |
| Farmer 1997 | 1991 - 1995 | 85 | 3.1‡ (2.1 - 4.5) | 5.0‡ (3.7 - 6.5) | - |
| Todd 1999 | 1992 - 1997 | 99 | 1 (reference) | 1.4 (0.7 - 2.8) | - |
| Bloemenkamp 1999 | 1994 - 1998 | 185 | 3.7 (1.9 - 7.2) | 5.6 (not given) | - |
| Parkin 2000 | 1990 - 1998 | 26 | 5.1 (1.2 - 21.4) | 14.9 (3.5 - 64.3) | - |
| Lidegaard 2002 | 1994 - 1998 | 987 | 2.9 (2.2 - 3.8) | 4.0 (3.2 - 4.9) | - |
| Dinger 2007 | 2000 - 2004 | 118 | 1 (reference) | 1.3 (NA) | 1.0 (0.6 - 1.8) |
| Vlieg 2009 | 1999 - 2004 | 1524 | 3.6 (2.9 - 4.6) | 7.3 (5.3 - 10.0)/5.6 (3.7 - 8.4) | 6.3 (2.9 - 13.7) |
| Lidegaard 2009 | 1995 - 2005 | 4213 | 2.0 (1.8 - 2.3) | 3.6 (3.3 - 3.8) | 4.0 (3.3 - 4.9) |
| Dinger 2010 | 2002 - 2008 | 680 | 1 (reference) | NA | 1.0 (0.6 - 1.8) |
| Parkin 2011 | 2002 - 2009 | 61 | 1 (reference) | NA | 2.7 (1.5 - 4.7) |
| Jick 2011 | 2002 - 2008 | 186 | 1 (reference) | NA | 2.8 (2.1 - 3.8) |
| Lidegaard 2011 | 2001 - 2009 | 4246 | 2.2 (1.7 - 2.8) | 4.2 (3.6 - 4.9) | 4.5 (3.9 - 5.1) |
| Confirmed only | 2001 - 2009 | 2707 | 2.9 (2.2 - 3.8) | 6.8 (5.7 - 8.1) | 6.3 (5.4 - 7.5) |
| FDA Kaiser 2011 | 2001 - 2007 | 625 | 1 (reference) | NA | 1.5 (1.2 - 1.9) |
| Gronich 2011 | 2002 - 2008 | 518 | 1 (reference) | 1.4 (0.9 - 2.1) | 1.7 (1.0 - 2.7) |
| Lidegaard 2012 | 2001 - 2010 | 5287 | 3.2 (2.7 - 3.8) | 6.5 (4.7 - 8.9)* | NA |
| Dinger 2014 | 2005 - 2010 | 162 | 1 (reference) | NA | 0.8 (0.5 - 1.6) |

‡ Absolute risk per 10,000 years.

* Vaginal ring with the third-generation progestin etonogestrel.

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Atherosclerosis and Cardiovascular Events

Noting that previous studies had demonstrated women on oral contraceptives (OC) faced a fourfold increased risk of heart attack (Hennekens 1977; Vessey 1976; Beral 1976), researchers in 1982 set out to understand the pathogenesis of vascular disease related to COCs. They found that combination oral contraceptives (COC) caused “greater cell proliferation and incorporation...in both human arterial smooth muscle cells and dermal fibroblasts.” Smooth muscle cell proliferation is an integral feature of all atherosclerotic lesions (Bagdade 1982).

In 2007, a presentation at the American Heart Association meeting described a study of 1,301 Belgian women, which showed that women had a 20 to 30 percent increase of plaque for every decade on COCs (Zoler 2007). They noted that active OC users had elevated C-reactive protein levels, three times higher than non-users. C-reactive protein is a biomarker for many inflammation-related arterial (and autoimmune) diseases, which was recently the subject of another presentation (Rietzschel 2018).

They evaluated the carotid and femoral pulse wave velocity (PWV) and found the average PWV among non-users was 6.6 m/sec, while the average among current OC users was 6.75 m/sec. The blood pressure of current OC users was also significantly higher (4.3/2.3 mm Hg higher than non-users) (Zoler 2007). Lead investigator Dr. Ernst Rietzschel said this study “changes our thinking about oral contraceptives just causing an increased thrombotic risk. Instead, it appears as though OC use may also cause long-term structural damage to the vasculature.” These findings were supported by an evaluation of large artery stiffness in the ENIGMA study (Hickson 2011) although other smaller studies have shown conflicting data (Yu 2014, Priest 2018).

A study of homocysteine and nitric oxide levels compared 50 healthy women with normal menstrual cycles as a control group and 50 healthy women receiving oral contraceptive pills for at least three menstrual cycles (Fallah 2012). They noted that after 3 months of treatment, homocysteine levels were significantly increased ($P = 0.027$), and there was a significant and considerable decrease ($P = 0.048$) in NO concentration of oral contraceptive pill (OCP) consumers. Another study evaluated the effect of COCs on homocysteine and C-reactive protein levels in women (Norouzi 2011). This observational cross-sectional analysis included 90 healthy, non-obese women (mean age 25 years). Forty-five healthy women on OCP and 45 healthy controls were studied. COC users had a minimum of 3 cycles on COCs. The results showed that the homocysteine (13.268 ± 3.475 vs. 7.288 ± 2.621 $\mu\text{mol/L}$) and CRP (5863.0 ± 1349.5 vs. 1138.3 ± 691.12 ng/ml) levels were significantly higher in women receiving OCP in comparison with the control group ($p=0.027$ and $p<0.001$, respectively). Similarly, a cross-sectional study, in 2011-2012, evaluated 60 healthy premenopausal women (30 cases of COC consumers and 30 controls as nonconsumers), aged between 25 and 45 years who were current users for at least a 3-year period. They evaluated brachial artery endothelial function (using flow-mediated dilatation (FMD)) and common carotid artery intima-media thickness (Heidarzadeh 2014). They noted that there was a significant FMD% difference between 2 groups of cases and controls: 11 ± 3.53 versus 15.80 ± 9.22 ($P = 0.01$). In addition, a significant mean CCA-IMT thickness difference was detected: 0.53 ± 0.07 versus 0.44 ± 0.08 ($P = 0.00$). Although these results were not significant after multiple regression analysis, the authors noted that their results were in favor of early atherosclerotic changes in prolonged users of COCs.

The Danish Heart Association released the results of a 15-year historic cohort study looking at thrombotic stroke and myocardial infarction, which observed over 1.6 million women. The results demonstrated that women taking COCs with ethinyl estradiol at a dose of 20 μg had a risk of arterial thrombosis that was 0.9 to 1.7 times higher than non-users, while those taking a dose of 30 to 40 μg had a 1.3 to 2.3 higher risk (Lidegaard 2012). The risk of thrombotic stroke appeared to be independent of duration of use, while the risk for myocardial infarction increased with duration of use (Table 15).

Together, these studies suggest that protracted use of COCs can induce atherosclerotic changes independent of any pro-thrombotic effect. These changes may contribute to the increase in thrombotic stroke and myocardial infarction seen in COC users.

Table 15 – Relative Risk of Thrombotic Stroke and Myocardial Infarction among Users of Selected Types of Combined Oral Contraception with Ethinyl Estradiol at a Dose of 30 to 40 µg, as Compared with Nonusers, According to Duration of Use (from Lidegaard 2012).

| Duration of use | No. of person-yrs. | Thrombotic Stroke | | Myocardial infarction | |
|-----------------|--------------------|-------------------|------------------------|-----------------------|------------------------|
| | | No. of events | Relative Risk (95% CI) | No. of events | Relative Risk (95% CI) |
| <1 year | 987,564 | 213 | 1.90 (1.64–2.20) | 86 | 1.85 (1.48–2.31) |
| 1-4 years | 992,825 | 194 | 1.55 (1.33–1.80) | 108 | 1.99 (1.63–2.43) |
| >4 years | 399,461 | 173 | 1.93 (1.65–2.26) | 91 | 2.11 (1.70–2.62) |

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Conclusion

Hormonal agents have a variety of effects on various organs and organ systems which may result in a deleterious impact on women's health. The data reviewed above reflect a vast body of information which has come to light since the introduction of these agents as contraceptives over 50 years ago. While the information for patients and prescribers currently reflects many of the known side effects, others have come to light which are not adequately represented in the current prescribing information. These should be added and made obvious to patients. In one instance, that of venous thromboembolism, while the warning information is present, it is phrased in a misleading manner which misleads the patients into drawing the incorrect conclusion regarding the risks. In addition, one agent (DMPA) appears to convey a specific risk for HIV transmission which is not shared by other agents. DMPA should be considered for revoking of marketing authorization and removed from the market. The risks of depression, mood disorders, and suicide have not been adequately emphasized.

We further encourage the Agency to require the manufacturers of these agents to widely publicize these additional risks. Many millions of women are currently receiving COCs and POCs. Many millions more have been exposed to these agents at some point in their lives. They should receive updated information regarding risks which have not been conveyed, or not adequately conveyed, in the past. All women who have been exposed to COCs or POCs should be informed so that they can take this information into account as they may encounter some of these adverse effects in some cases many years after cessation of use.

Environmental Impact

Based on data from the Guttmacher Institute, a conservative estimate of 11 million women aged 15-44 in the US take some form of hormonal contraceptive each day¹⁵⁵. A 2015 study reports that about 21 percent of women of reproductive years are using some form of hormonal contraceptive, which equates to about 13 million women (Daniels 2015). This has resulted in a significant increase in the release of synthetic progestagens (such as levonorgestrel) and synthetic estrogens (such as ethinylestradiol [EE2]) into the aquatic environment via wastewater treatment plant discharges (Besse 2009, King 2016). EE2 is metabolized in the liver undergoing first pass metabolism, but ~6% of the administered dose appears as untransformed EE2 in the urine and ~9% in the feces (Stanczyk 2013). As noted by King (King 2016), even at low concentrations, these compounds can act as potent endocrine disruptors, affecting the growth, development, and reproduction of exposed aquatic organisms (Tyler 1998, Larsson 1999). EE2 is one of the most studied synthetic hormones in aquatic environments, for which assessments of environmental concentrations and the quantification of endocrine-related effects have been documented in a range of aquatic species (Purdom 1994, Jobling 1998, Kirby 2004, Jobling 2006). In fact, the numerous studies on the effects of EE2 on aquatic organisms have led to the derivation of a reliable predicted no-effect concentration of 0.1 ng/L for EE2 (Caldwell 2012).

In 1993, the first publication appeared which brought attention to the issue of synthetic chemicals mimicking natural estrogen in the environment (Sharpe 1993). The study pointed to environmental pollutants, which were having a deleterious effect on male fetuses in utero – endocrine disruptors like polychlorinated biphenyls, detergents, dioxins, and hormonal contraceptives. In 1995, another paper (Sumpter 1995) noted that male fish in 28 rivers across Britain were being “feminized” by pollutants. In 2002, a paper was published that focused specifically on the effects of endocrine-disrupting chemicals in the environment (Jobling 2002). They demonstrated reduced fertility in fish populations in areas downstream of effluent from sewage plants

¹⁵⁵ <https://www.guttmacher.org/fact-sheet/contraceptive-use-united-states>.

located along tributaries of the Thames River. In 2007, the results of a seven-year Canadian lake study were published which examined the effects of EE2 (Kidd 2007). The researchers released a quantity of EE2 equivalent to what would come into the waterways via sewage from a city of 200,000 people. They witnessed an immediate feminization and transgendering of male fish, which resulted in the “near extinction” of the fathead minnow population (Kidd 2007). Although the minnow populations neared extinction, they rebounded as soon as the researchers stopped adding EE2 to the lake. A 2006 study from the United States Geological Survey on smallmouth bass in the Shenandoah and Monocacy Rivers found that more than 80-percent of all the male bass living in these waterways were growing eggs in their testes¹⁵⁶.

A study was carried out of fish populations relative to the sewage treatment plants of three major Colorado cities: Denver, Boulder, and Colorado Springs (Woodling 2006). At each municipality, they set up a location just upstream from where the effluent was released, and another just downstream. The fish in the upstream locations enjoyed a balanced 1:1 female-to-male sex ratio. Downstream there were five female fish for every one male, and twenty percent of the reduced male population demonstrated intersex characteristics, such as eggs in their testes and the presence of vitellogenin, an egg yolk protein normally found only in fertile females. The consequences also appeared to ascend up the food chain in a measurable way, specifically with the feminization of trout, mink frogs and green frogs (Parke 2009). Both the predicted and the measured concentrations of EE2 in the US, including effluent of waste water treatment plants, surface water, or ground water, exceeds the predicted no-effect concentrations on fish populations (Kostich 2013).

Environmental factors have been implicated in declining fertility rates (Skakkebaek 2016). A 2017 study out of Hebrew University and Mount Sinai Medical School found that sperm counts in human men have dropped by more than 50 percent since 1973 (Levine 2017). While it has been noted that environmental exposure to individual steroidal estrogens, as well as their mixtures, are unlikely to dramatically affect endocrine signaling in humans, it is not clear whether more subtle effects are possible (Kostich 2013). More recently, environmental effects of levonorgestrel have been postulated (King 2016) but there is less hard data.

There is a clear effect of environmental EE2 on fish populations as well as species higher in the food chain such as frogs. An effect on humans is also possible.

Environmental Impact References

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¹⁵⁶ https://dep.wv.gov/WWE/watershed/wqmonitoring/Documents/Potomac-Intersex/USGS_FishHealthReproductivelssuesPotomac_2006.pdf.

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Economic Impact

For the diseases noted below, in some cases we have calculated the estimated economic impact taking into account those who are currently using COCs and those who have ever used COCs. According to the CDC¹⁵⁷ 15.9% of women aged 15–44 in the US use “the pill.” There are 61 million US women of reproductive age (15–44)¹⁵⁸. This yields 9,699,000 women in the USA currently on COCs. Note that this is a low estimate as it does not include women using intravaginal and transdermal formulations and is lower than the estimate by Daniels (Daniels 2015).

According to the National Survey of Family Growth¹⁵⁹, 79.3% of women surveyed from 2011–2015 have ever used “the pill.” This is down from 81.9% in the 2006–2010 survey and 82.3% in the 2002 survey. The lower number for “ever use” of 79.3% is used in subsequent calculations. According to the 2010 census (Howden 2011), there were 156,964,212 women in the US, of whom 24% were under 18 years of age. Thus, there were 119,292,801 women 18 years of age or older. This implies that $119,292,801 \times 0.793 = 94,599,191$ women in the USA have ever used the pill. As noted above, this does not include women using intravaginal and transdermal formulations.

The numbers 9,699,000 for current use and 94,599,191 for ever use of COCs were used in some of these calculations. In other cases, the census data for specific age groups was used if they were the groups most likely to be impacted by current or recent use of COCs.

For progesterone-only contraceptives (POCs), the National Survey of Family Growth¹⁶⁰, notes that 25.4% of women aged 15–44 in 2011–2015 have ever used “3-month injectable (Depo-Provera™).” This is up from 23.2% in 2006–2010 and 16.8% in 2002. For a conservative estimate, we will use the lowest of these numbers (16.8% or 20,041,191 women) who have ever used POCs. This would not include POCs administered by other routes and is thus a conservative estimate.

HIV Costs

According to the CDC¹⁶¹, an estimated 255,900 women were living with HIV at the end of 2014. Of these it is estimated 87% were via sexual contact (this proportion was relatively stable from 2011–2016; CDC HIV Surveillance Table 1a). Annual medical cost estimates for HIV-infected persons, adjusted for age, sex, race/ethnicity, and transmission risk group, were from the HIV Research Network (range \$1,854–\$4,545/month) and for HIV-uninfected persons were from the Medical Expenditure Panel Survey (range \$73–\$628/month) (Schackman 2015). Using this information along with the prevalence of DMPA use of 16.8%, this suggests an annual cost of treatment for HIV infection due to DMPA use of ~\$157–573 million (Table 16).

¹⁵⁷ <https://www.cdc.gov/nchs/fastats/contraceptive.htm>.

¹⁵⁸ <https://www.cdc.gov/nchs/data/nhsr/nhsr086.pdf>.

¹⁵⁹ https://www.cdc.gov/nchs/nsfg/key_statistics/c.htm#everused.

¹⁶⁰ https://www.cdc.gov/nchs/nsfg/key_statistics/c.htm#everused.

¹⁶¹ <https://www.cdc.gov/hiv/group/gender/women/index.html>.

Table 16 – Estimated Economic Impact of DMPA due to Increased Prevalence of HIV Infection

| | |
|---|---------------|
| Women with HIV | 255,900 |
| Sexual transmission | 87% |
| Cases due to sexual transmission | 222,633 |
| Ever use of DMPA | 16.80% |
| Women with HIV with DMPA use | 37,402 |
| RR of HIV with DMPA use | 1.4 |
| Adjusted estimate → | 26,716 |
| Excess cases → | 10,686 |
| Highest estimated individual annual costs → | \$53,664 |
| Lowest estimated individual annual costs → | \$14,712 |
| Highest estimated total annual costs → | \$573,474,111 |
| Lowest estimated total annual costs → | \$157,218,081 |

Breast Cancer

A recent study in the US (Blumen 2016) notes, “The average costs per patient allowed by the insurance company in the year after diagnosis were \$60,637, \$82,121, \$129,387, and \$134,682 for disease stage 0, I/II, III, and IV, respectively. The average costs allowed per patient in the 24 months after the index diagnosis were \$71,909, \$97,066, \$159,442, and \$182,655 for disease stage 0, I/II, III, and IV, respectively.” For all patients, they note that the average cost for the first 12 months following diagnosis is \$85,772, and for the second 12 months is \$22,127 with a total of \$103,735 for the 24 months following diagnosis. For these calculations we will use the first-year costs to estimate costs for incident cases among current users of COCs and will use the second-year cost to approximate the average annual cost of care for a patient diagnosed with breast cancer. According to the NIH SEER statistics¹⁶², the incidence of breast cancer is 126.0 per 100,000 person-years. Approximately 12.4 percent of women will be diagnosed with female breast cancer at some point during their lifetime. According to the best epidemiology studies noted in Table 3 (Mørch 2017; Heikkinen 2016, Lund 2007), and the best meta-analysis in Table 5 (Kahlenborn 2006) the relative risk of ever use of COCs for the development of breast cancer is 1.19–1.37. Based on this information, the estimated increase in cost from use of COCs due to incident cases of breast cancer is between \$199 million and \$387 million (Table 17).

Table 17 – Estimated Economic Impact of COCs due to Increased Incidence of Breast Cancer

| | | | | |
|--|--------------------|---------------|---------|---------|
| Women of reproductive age | Number on the pill | Incidence | | |
| 61,000,000 | 9,699,000 | 0.00126 | | |
| Estimated women on the pill at risk → | | 12,221 | | |
| Adjusted estimate of cases → | | 14,543 | 1.19 | Low RR |
| Adjusted estimate of cases → | | 16,742 | 1.37 | High RR |
| Excess cases → | | 2,322 | Low RR | |
| Excess cases → | | 4,522 | High RR | |
| Annual cost per patient of breast cancer → | | \$85,772 | | |
| Estimated annual costs → | | \$199,157,489 | Low RR | |
| Estimated annual costs → | | \$387,833,005 | High RR | |

¹⁶² <https://seer.cancer.gov/statfacts/html/breast.html>.

To evaluate the impact of “ever use” of COCs on prevalent breast cancer, we noted that the best meta-analysis (Kahlenborn 2006) showed a 1.19 odds ratio of breast cancer with COCs. According to the SEER statistics, there are currently 3,418,124 prevalent cases of breast cancer in the USA. The estimated increase in cost from treatment of the excess cases of breast cancer is estimated to be ~\$9.6 billion annually (Table 18).

Table 18 – Estimated Economic Impact of COCs due to Increased Prevalence of Breast Cancer

| Prevalent cases of breast cancer | Ever use of COCs | Breast cancer ever users | | |
|--|------------------|--------------------------|------|-----------------|
| 3,418,124 | 79.3% | 2,710,572 | | |
| Adjusted estimate of cases if no use of COCs → | | 2,277,792 | 1.19 | RR |
| Excess cases → | | 432,780 | | |
| Annual cost per patient of breast cancer → | | | | \$22,127 |
| Estimated total costs → | | | | \$9,576,133,158 |

Cervical Cancer

A recent study in Canada (Pendrith 2016) on the costs of invasive cervical cancer treatment noted: “The mean overall medical care cost was \$39,187 [standard error (se): \$1,327] in the 1st year after diagnosis. ... At 5 years after diagnosis, the mean overall unadjusted cost was \$63,131 (se: \$3,131), and the cost adjusted for censoring was \$68,745 (se: \$2,963).” For these calculations we will assume a cost of \$39,187 annually for incident cases and \$13,749 (= \$68,745/5) annually for prevalent cases of invasive cervical cancer. According to the NIH SEER statistics¹⁶³, the incidence of invasive cervical cancer is 7.4 per 100,000 person-years. According to the American Cancer Society¹⁶⁴, it is estimated that 13,170 women will be diagnosed with invasive cervical cancer in the USA in 2019. In 2015, there were an estimated 257,524 women living with invasive cervical cancer in the United States. According to the best epidemiology studies noted in Table 6 (Roura 2016) the relative risk of ever use of COCs for the development of invasive cervical cancer is 1.6 and the RR for current use is 2.2. Based on this information, the estimated increase in cost from use of COCs due to incident cases of cervical cancer is ~\$33 million (Table 19).

Table 19 – Estimated Economic Impact of COCs due to Increased Incidence of Cervical Cancer

| Women of reproductive age | Number on the pill | Incidence | | |
|--|--------------------|--------------|-----|----|
| 61,000,000 | 9,699,000 | 0.000074 | | |
| Estimated women on the pill at risk → | | 718 | | |
| Adjusted estimate of cases → | | 1,579 | 2.2 | RR |
| Excess cases → | | 861 | | |
| Annual cost per patient of cervical cancer → | | \$39,187 | | |
| Estimated annual costs → | | \$33,750,635 | | |

To evaluate the impact of “ever use” of COCs on prevalent cervical cancer, we noted that the best study (Roura 2016) showed a 1.6 relative risk of cervical cancer with COCs. According to the SEER statistics, there are currently 257,524 prevalent cases of cervical cancer in the USA. The estimated increase in cost from treatment of the excess cases of cervical cancer is estimated to be ~\$1 billion annually (Table 20).

¹⁶³ <https://seer.cancer.gov/statfacts/html/cervix.html>.

¹⁶⁴ <https://www.cancer.org/cancer/cervical-cancer/about/key-statistics.html>.

Table 20 – Estimated Economic Impact of COCs due to Increased Prevalence of Cervical Cancer

| | | | | |
|--|-------------------------|----------------------------|-----------------|----|
| Prevalent cases of breast cancer | Ever use of COCs | Cervical cancer ever users | | |
| 3,418,124 | 79.3% | 257,524 | | |
| Adjusted estimate of cases if no use of COCs → | | 204,217 | 1.6 | RR |
| | Excess cases → | 76,581 | | |
| Annual cost per patient of cervical cancer → | | | \$13,749 | |
| | Estimated total costs → | | \$1,052,914,912 | |

Crohn’s Disease

A recent study in the US (Rao 2018) estimated the 5-year cost of the treatment of Crohn’s disease as \$116,838 per patient (interquartile range of \$45,643–\$240,398; annual cost \$23,368). This was higher with worsening disease activity. According to the Centers for Disease Control (CDC), the incidence of Crohn’s disease is 3.1 to 14.6 cases per 100,000 person-years¹⁶⁵. According to the best epidemiology studies noted in Table 7 (Khalili 2013; García Rodríguez 2005), and the best meta-analysis (Cornish 2008), the relative risk of current COC use is 1.46–2.82 for the development of Crohn’s disease. Based on this information, the estimated increase in cost just from treatment of the excess cases of Crohn’s disease, only looking at current use and not past use of COCs, is between \$3 million and \$60 million annually (Table 21).

Table 21 – Estimated Economic Impact of COCs due to Increased Incidence of Crohn’s Disease

| | | | | | |
|---------------------------|--|---------------|----------------|---------|---------|
| Women of reproductive age | Number on the pill | Low incidence | High incidence | | |
| 61,000,000 | 9,699,000 | 0.000031 | 0.000146 | | |
| | Estimated women on the pill at risk → | 301 | 1,416 | | |
| | Adjusted estimate → | 439 | 2,067 | 1.46 | Low RR |
| | Adjusted estimate → | 848 | 3,993 | 2.82 | High RR |
| | Excess cases → | 138 | 651 | Low RR | |
| | Excess cases → | 547 | 2,577 | High RR | |
| | Annual cost per patient of Crohn’s disease → | \$23,368 | | | |
| | Estimated annual costs → | \$3,231,920 | \$15,221,300 | Low RR | |
| | Estimated annual costs → | \$12,787,162 | \$60,223,406 | High RR | |

To evaluate the impact of “ever use” of COCs, we noted that the best cohort study (Khalili 2013) and meta-analysis (Cornish 2008) showed a 1.43 and 1.44 relative risk of Crohn’s disease. According to the Centers for Disease Control (CDC), the prevalence of Crohn’s disease in adults is 201 cases per 100,000 person-years¹⁶⁶. Taking the lower number of 1.43, the estimated increase in cost from treatment of the excess cases of Crohn’s disease due to COC use is approximately \$1.9 billion annually (Table 22).

Table 22 – Estimated Economic Impact of COCs due to Increased Prevalence of Crohn’s Disease

| | | | |
|---------------------------|------------------|------------|--|
| Women ≥ 18 in 2010 Census | Ever use of COCs | Prevalence | |
|---------------------------|------------------|------------|--|

¹⁶⁵ <https://www.cdc.gov/ibd/IBD-epidemiology.htm>.

¹⁶⁶ <https://www.cdc.gov/ibd/IBD-epidemiology.htm>.

| | | | | |
|--|------------|-----------------|------|----|
| 119,292,801 | 94,599,191 | 0.000201 | | |
| Estimated women on the pill at risk → | | 190,144 | | |
| Adjusted estimate → | | 271,906 | 1.44 | RR |
| Excess cases → | | 81,762 | 1.44 | RR |
| Annual cost per patient of Crohn's disease → | | \$23,368 | | |
| Estimated total costs → | | \$1,910,583,605 | 1.44 | RR |

Ulcerative Colitis

A recent study in the US (Cohen 2015) noted that compared with controls, patients with UC had higher adjusted total direct (\$15,548 vs \$4812) and indirect costs (\$4125 vs \$1961) annually. This implies a total annual increase in cost of ~\$12,900 for UC. This was higher with worsening disease activity. According to the Centers for Disease Control (CDC), the incidence of UC is 2.2 to 14.3 cases per 100,000 person-years¹⁶⁷. According to the best epidemiology studies noted in Table 8 (Khalili 2013; García Rodríguez 2005), and the best meta-analysis (Cornish 2008) the relative risk of current COC use 1.22–1.58 for the development of UC. Based on this information, the estimated increase in cost just from treatment of the excess cases of UC, only looking at current use and not past use of COCs is between \$605,000 and \$10 million per year (Table 23).

Table 23 – Estimated Economic Impact of COCs due to Increased Incidence of Ulcerative Colitis

| Women of reproductive age | Number on the pill | Low incidence | High incidence | | |
|---|--------------------|---------------|----------------|---------|---------|
| 61,000,000 | 9,699,000 | 0.000022 | 0.000143 | | |
| Estimated women on the pill at risk → | | 213 | 1,387 | | |
| Adjusted estimate → | | 260 | 1,692 | 1.22 | Low RR |
| Adjusted estimate → | | 337 | 2,191 | 1.58 | High RR |
| Excess cases → | | 47 | 305 | Low RR | |
| Excess cases → | | 124 | 804 | High RR | |
| Annual cost per patient of ulcerative colitis → | | \$12,900 | | | |
| Estimated annual costs → | | \$605,567 | \$3,936,184 | Low RR | |
| Estimated annual costs → | | \$1,596,494 | \$10,377,212 | High RR | |

To evaluate the impact of “ever use” of COCs, we noted that the best cohort study (Khalili 2013) showed a 1.18 relative risk of UC. The estimated increase in cost of the excess cases of UC due to use of COCs is approximately \$522 million annually (Table 24).

Table 24 – Estimated Economic Impact of COCs due to Increased Prevalence of Ulcerative Colitis

| Women ≥ 18 in 2010 Census | Ever use of COCs | Prevalence | | |
|---|------------------|---------------|------|----|
| 119,292,801 | 94,599,191 | 0.000238 | | |
| Estimated women on the pill at risk → | | 225,146 | | |
| Adjusted estimate → | | 265,672 | 1.18 | RR |
| Excess cases → | | 40,526 | 1.18 | RR |
| Annual cost per patient of ulcerative colitis → | | \$12,900 | | |
| Estimated total costs → | | \$522,789,187 | 1.18 | RR |

¹⁶⁷ <https://www.cdc.gov/ibd/IBD-epidemiology.htm>.

Systemic Lupus Erythematosus

A recent study in the US (Chen 2015) noted that mean total health care costs were \$21,535 among all SLE patients over the 1-year study period. According to the Centers for Disease Control (CDC), the incidence of SLE is 6.5–10.6 cases per 100,000 women-years¹⁶⁸. In terms of prevalence, “A conservative estimate suggests a prevalence of 161,000 with definite SLE and 322,000 with definite or probable SLE.” According to the best epidemiology studies noted in Table 9 that evaluated current use of COCs (Bernier 2009), the relative risk of current COC use is 1.45 – 2.52 for the development of SLE. Based on this information, the estimated increase in cost just from treatment of the excess cases of SLE, only looking at current use and not past use of COCs, is \$6.1 million to \$33.6 million annually (Table 25).

Table 25 – Estimated Economic Impact of COCs due to Increased Incidence of Systemic Lupus Erythematosus.

| Women of reproductive age | Number on the pill | Low incidence | High incidence | | |
|---------------------------------------|--------------------|---------------|----------------|---------|---------|
| 61,000,000 | 9,699,000 | 0.000065 | 0.0001065 | | |
| Estimated women on the pill at risk → | | 630 | 1,028 | | |
| Adjusted estimate → | | 914 | 1,491 | 1.45 | Low RR |
| Adjusted estimate → | | 1,589 | 2,591 | 2.52 | High RR |
| Excess cases → | | 284 | 463 | Low RR | |
| Excess cases → | | 958 | 1,563 | High RR | |
| Annual cost per patient of SLE → | | \$21,535 | | | |
| Estimated annual costs → | | \$6,109,388 | \$9,963,002 | Low RR | |
| Estimated annual costs → | | \$20,636,155 | \$33,652,807 | High RR | |

To evaluate the impact of “ever use” of COCs, we noted that the best cohort studies (Costenbader 2007; Bernier 2009) showed a relative risk of SLE 1.19–2.3. The estimated increase in cost of the excess cases of SLE due to use of COCs is approximately \$439 million–\$1.55 billion annually (Table 26).

Table 26 – Estimated Economic Impact of COCs due to Increased Prevalence of Systemic Lupus Erythematosus.

| Women ≥ 18 in 2010 Census | Ever use of COCs | Prevalence | | | |
|---------------------------------------|------------------|-----------------|------|---------|--|
| 119,292,801 | 94,599,191 | 161,000 | | | |
| Estimated women on the pill at risk → | | 127,673 | | | |
| Adjusted estimate → | | 107,288 | 1.19 | Low RR | |
| Adjusted estimate → | | 55,510 | 2.3 | High RR | |
| Excess cases → | | 20,385 | 1.19 | Low RR | |
| Excess cases → | | 72,163 | 2.3 | High RR | |
| Annual cost per patient of SLE → | | \$21,535 | | | |
| Estimated total costs → | | \$438,985,908 | 1.19 | Low RR | |
| Estimated total costs → | | \$1,554,030,205 | 2.3 | High RR | |

Depression

The most reliable study (Skovlund 2016) indicated a 1.1 RR for depression with COCs and a 1.2 RR with POCs. This study evaluated women aged 15-34 and then followed them for a mean of 5 years. According to the information from Brody (Brody 2018), the prevalence of depression in women aged 20-39 is 10.1%. An analysis

¹⁶⁸ <https://www.cdc.gov/lupus/facts/detailed.html>.

of medical claims conducted by insurer Blue Cross Blue Shield (Blue Cross Blue Shield 2018) found that “in 2016, Blue Cross plans spent \$10,673 on those diagnosed with ‘major depression’ compared to \$4,283 on those without a depression diagnosis.” With this information, and noting from the census data (Howden 2011) that there are ~52 million women aged 15-39, we calculate that the excess annual cost of depression attributable to COCs is ~\$2.4 billion (Table 27) and from POCs is ~\$937 million (Table 28).

Table 27 – Estimated Economic Impact of COCs due to Increased Prevalence of Depression

| | |
|--------------------------------------|-----------------|
| Women aged 15-39 | 51,877,977 |
| Percent with depression | 10.1% |
| Women aged 15-39 with depression | 5,239,675.68 |
| Ever use of COCs | 79.30% |
| 15-39 y.o. COC users with depression | 4,155,063 |
| RR of depression with COC use | 1.1 |
| Adjusted estimate → | 3,777,330 |
| Excess cases → | 377,733 |
| Estimated individual annual costs → | \$6,390 |
| Estimated total annual costs → | \$2,413,713,761 |

Table 28 – Estimated Economic Impact of POCs due to Increased Prevalence of Depression

| | |
|--------------------------------------|---------------|
| Women aged 15-39 | 51,877,977 |
| Percent with depression | 10.1% |
| Women aged 15-39 with depression | 5,239,675.68 |
| Ever use of POCs | 16.80% |
| 15-39 y.o. COC users with depression | 880,266 |
| RR of depression with POC use | 1.2 |
| Adjusted estimate → | 733,555 |
| Excess cases → | 146,711 |
| Estimated individual annual costs → | \$6,390 |
| Estimated total annual costs → | \$937,482,772 |

Multiple Sclerosis

As the most rigorous cohort studies did not show an increase in the risk of developing multiple sclerosis a rigorous cost analysis was not performed. However, using the information from the best case-control study (Hellwig 2016), an increased odds ratio of 1.51 was noted. If this is assumed to be accurate, this can be used along with a study of total MS costs from 1997-2013 (Chen 2017). They noted that, “The total charges on managing MS range from \$161 million in 1997 to \$755 million in 2013.” Conservatively assuming steady costs since 2013, we can calculate that 79.3% of those costs were incurred by women who were “ever users” of

COCs. This yields \$598,715,000. If these women had not used COCs there would have been a proportionate reduction in costs of \$202,215,000 ($\$598,715,000 - (\$598,715,000/1.51)$).

Interstitial Cystitis

According to one recent paper (Tung 2017) on average, having interstitial cystitis was associated with \$7,223 higher total health care costs annually than not having IC. The prevalence of interstitial cystitis has been estimated at 2.7% using a high specificity definition (McLennan 2014) while another study in a managed care population (Clemens 2005) indicated (depending on the definition) a prevalence between 45 and 197 per 100,000 women. Using the most conservative estimate (Champaneria 2015) “ever use” of COCs is associated with an OR of 2.31 for interstitial cystitis. Assuming 61 million women of reproductive age, with a 79.3% of exposure to COCs, this suggests ~11,500 excess cases (using a prevalence of interstitial cystitis of 45/100,000) to ~50,500 (using a prevalence of interstitial cystitis of 197/100,000). This yields an annual cost of \$83–\$365 million (Table 29).

Table 29 – Estimated Annual Economic Impact of COCs due to Increased Prevalence of Interstitial Cystitis

| | |
|---|---------------|
| Low prevalence of interstitial cystitis | 0.00045 |
| High prevalence of interstitial cystitis | 0.00197 |
| Women of reproductive age | 61,000,000 |
| Number with ever use of the pill | 48,373,000 |
| # of Women with interstitial cystitis low prevalence | 21,768 |
| # of Women with interstitial cystitis high prevalence | 95,295 |
| OR | 2.13 |
| Excess cases of interstitial cystitis low prevalence | 11,548 |
| Excess cases of interstitial cystitis high prevalence | 50,555 |
| Annual cost | \$7,223 |
| Annual cost of interstitial cystitis low prevalence | \$83,412,664 |
| Annual cost of interstitial cystitis high prevalence | \$365,162,106 |

Osteoporotic Bone Fracture Risk

According to a recent review (Ballane 2017), in North America the incidence of osteoporotic vertebral fractures is 837 to 1,083 cases per 100,000 women per year (mean of 960 per 100,000 per year) as standardized to 2015. The annual excess cost of care for women with osteoporotic vertebral fractures was estimated to be \$11,655 per year (Kilgore 2009). Using the most relevant relative risk of 1.07 (Barad 2005), this implies an annual cost of ~\$308 million dollars in the US from COC use (Table 30).

Table 30 – Estimated Economic Impact of COCs due to Increased Annual Incidence of Vertebral Fractures

| | | | | |
|---|------------------|---|------|----|
| Women ≥ 50 in 2010 Census | Ever use of COCs | Incidence of osteoporotic vertebral fractures | | |
| 53,151,456 | 42,149,105 | 0.0096 | | |
| Estimated women on the pill with Fx → | | 404,631 | | |
| Adjusted estimate → | | 378,160 | 1.07 | RR |
| Excess cases → | | 26,471 | | |
| Annual cost per patient of osteoporotic vertebral fractures → | | \$11,655 | | |

| | |
|--------------------------------|---------------|
| Estimated total annual costs → | \$308,521,992 |
|--------------------------------|---------------|

The best cohort study on fracture risk with progesterone-only contraceptives (POCs) showed a RR of 1.51 for ever use of DMPA (Lanza 2013), the most widely used POC. Assuming 16.8% of women have used POCs this yields an annual cost of ~\$290 million dollars in the US from POC use (Table 31).

Table 31 – Estimated Economic Impact of POCs due to Increased Annual Incidence of Vertebral Fractures

| | | | | |
|---|------------------|---|------|----|
| Women ≥ 50 in 2010 Census | Ever use of POCs | Incidence of osteoporotic vertebral fractures | | |
| 53,151,456 | 8,929,445 | 0.0096 | | |
| Estimated women on the pill with Fx → | | 85,723 | | |
| Adjusted estimate → | | 60,796 | 1.41 | RR |
| Excess cases → | | 24,926 | | |
| Annual cost per patient of osteoporotic vertebral fractures → | | \$11,655 | | |
| Estimated total annual costs → | | \$290,517,770 | | |

Body Mass

The costs of the effects on body mass were not calculated, but these effects are contributory to atherosclerosis and cardiovascular events, which are discussed below.

Urogenital Effects

The medical and societal costs of the urogenital effects of hormonal contraceptives were not calculated as, although there are measurable costs, they are not felt to be significant.

Venous Thromboembolism, Atherosclerosis and Cardiovascular Disease

About 1 in every 4 female deaths is due to heart disease; it is the leading cause of death for women in the U.S.¹⁶⁹ A review of recent population studies revealed that the overall prevalence of Peripheral Arterial Disease (PAD) for women is 15.6% (compared to 13.4% for men).¹⁷⁰ In 2008, coronary heart disease was prevalent in 7.5 million women.¹⁷¹ The total mean direct medical costs for cardiovascular disease (CVD) is \$18,953 annually (Nichols 2010). Using the median relative risk of the most popular birth control brands, the RR is 1.8 (Table 32).

¹⁶⁹ https://www.cdc.gov/dhds/data_statistics/fact_sheets/fs_women_heart.htm.

¹⁷⁰ https://www.medscape.org/viewarticle/711179_2.

¹⁷¹ <https://www.healthline.com/health/heart-disease/women-statistics-facts#1>.

Table 32 – Estimated Economic Impact of COCs due to Increased Incidence of Cardiovascular Disease

| | | | | |
|--|------------------|-------------------|-----|----|
| Coronary heart disease | Ever use of COCs | CHD in ever users | | |
| 7,500,000 | 79.3% | 5,947,500 | | |
| Adjusted estimate of cases if no use of COCs → | | 3,304,167 | 1.8 | RR |
| Excess cases → | | 2,643,333 | | |
| Annual cost per patient for CVD → | | \$18,953 | | |
| Estimated total costs → | | \$50,099,090,349 | | |

A more conservative estimate would assume that the increased risk for cardiovascular disease is limited to women aged 15–49 years, which was the group studied by Lidegaard (Lidegaard 2012). According to the US Census in 2010, population is broken down by age group (Howden 2011). The rate of cardiovascular events is similarly broken down by Lidegaard (Lidegaard 2012). Thus, the number of cases by age group is shown in Table 33.

Table 33 – Cardiovascular Events in Women by Age Group

| Census data | | Events per 100,000 person-years (Lidegaard 2012) | | Events per year | |
|---------------------------------|-----------------|---|--------|-------------------------|----------|
| Age group | Number of women | Myocardial infarction | Stroke | Myocardial infarction # | Stroke # |
| 15 to 19 years | 10,736,677 | 0.4 | 3.4 | 43 | 365 |
| 20 to 24 years | 10,571,823 | 0.7 | 5.6 | 74 | 592 |
| 25 to 29 years | 10,466,258 | 2.2 | 10.5 | 230 | 1,099 |
| 30 to 34 years | 9,965,599 | 5 | 15.4 | 498 | 1,535 |
| 35 to 39 years | 10,137,620 | 12.2 | 23.3 | 1,237 | 2,362 |
| 40 to 44 years | 10,496,987 | 25.4 | 39.2 | 2,666 | 4,115 |
| 45 to 49 years | 11,499,506 | 38.2 | 64.4 | 4,393 | 7,406 |
| Total number of events per year | | | | 9,141 | 17,473 |

Using these estimates, with the annual cost of care for cardiovascular disease and the relative risk noted above, this calculates to ~\$61 million in excess costs for myocardial infarctions and ~\$117 million in excess costs for strokes (Table 34).

Table 34 – Cost of Cardiovascular Events in Women Attributable to COC use.

| | Myocardial infarction | Stroke | | |
|---------------------------------|-----------------------|---------------|-----|-------------|
| Total events per year | 9,141 | 17,473 | | |
| Ever use of the pill | 79.30% | | | |
| # with events on COCs | 7,249 | 13,856 | 1.8 | RR Ever Use |
| Adjusted estimate → | 4,027 | 7,697.96 | | |
| Excess cases → | 3,222 | 6,158 | | |
| Estimated annual costs → | \$18,953 | \$18,953 | | |
| Estimated Excess annual costs → | \$61,062,935 | \$116,719,504 | | |

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Certification

We certify that this petition contains all relevant information, including any that may be unfavorable to the petition, that we were able to obtain.

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